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# Pedestrian-Vehicle Interaction in a CAV Environment: Explanatory Metrics 

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CENTER FOR CONNECTED AND AUTOMATED TRANSPORTATION

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# Pedestrian-Vehicle Interaction in a CAV Environment: Explanatory Metrics 

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## LIST OF ABBREVIATIONS AND ACRONYMS

| GroupSize | The number of pedestrians in the curb area, including the subject pedestrian |
| :---: | :---: |
| AgeRange | Estimated age range (yrs) for subject pedestrian(s) - 1(0-10), 2(10-30), 3(30-50), 4(>50) |
| Sex | Sex of subject pedestrian |
| Hesitation | Does the pedestrian slow down or wait at curb? |
| Distraction | Does a pedestrian approach and/or cross while using a cellphone or talking? |
| FlowWith | The number of pedestrians already crossing in the crosswalk in the same direction when subject pedestrian arrives at curb area |
| FlowAgainst | The number of pedestrians already crossing in the crosswalk in the opposite direction when subject pedestrian arrives at curb area |
| FlowOn | Total number of pedestrians already crossing in the crosswalk when an interaction occurs (FlowWith + FlowAgainst) |
| SameDirec | The number of pedestrians present in the curb area crossing in the same direction as the subject pedestrian |
| DiffDirec | The number of pedestrians present in a curb area with crossing direction opposite of the subject pedestrian |
| PedWait | Total number of pedestrians waiting in the curb areas when an interaction occurs (SameDirec + DiffDirec) |
| ApprSpeed | The approach speed of interacted vehicles when a pedestrian enters the curb area. (mph) |
| SlowsDown | Does a vehicle slow down or stop on the approach to the crosswalk when a pedestrian enters the curb area? |
| CloseFollow | Does the interacted vehicle have a close follower when an interaction occurs? |
| AdjVeh | Is a vehicle already present in the adjacent lane when a motorist begins to interact with a pedestrian? |
| Distance | The distance of interacted vehicle(s) to subject pedestrians when interaction begins. (in ft.) |
| NoF | Is pedestrian entering curb area on the near side or far side of the approaching vehicle's lane? |
| Pedestrian Outcomes | Cross: $\mathrm{Y}=1$; Wait/Yield: $\mathrm{Y}=0$ |
| Vehicle <br> Response | Level of vehicle deceleration when pedestrian enters crosswalk ( 3 = stops; 2 = slows down; $1=$ Does not slow down). |

## 1. INTRODUCTION

### 1.1. Background and Problem Statement

"State Law Yield to Pedestrian Within Crosswalk" signs (Figure 1(a)) are commonly used at unsignalized pedestrian crosswalks where pedestrian-motorist interaction frequently occurs. Pedestrians using crosswalks with "State Law Yield to Pedestrian Within Crosswalk" signs theoretically have priority over approaching vehicles. Nevertheless, observations confirm that confusion exists among pedestrians and motorists, because the sign's message is subject to varying interpretations. At certain times, a motorist stops and lets pedestrians standing at the curb cross the street, and sometimes, drivers fail to yield to pedestrians entering the crosswalk. In many cases, a non-verbal "negotiation" takes place between pedestrians and motorists, to determine who should proceed first.

Consequently, confusion between pedestrians and motorists leads to unnecessary delays for both pedestrians and vehicles and increase risks to pedestrian safety. An investigation of pedestrian crossing behavior and waiting behavior at such locations can be useful in developing policies and control strategies to enhance a pedestrian's perceived safety and improve the level of service (LOS) at unsignalized intersections.


Figure 1.Semi-Controlled Crosswalk
(a) Sign at Semi-Controlled Crosswalk. (MUTCD, 2009)
(b) One-way North University Street at Second Street, 2017 (Jon D. Fricker)
(c) Two-way North University Street at Second Street, 2018. (Google Maps)

### 1.2. Study Objective and Research Questions

Video recordings at crosswalks were created, making possible analyses that not only supplement gap acceptance methods to model pedestrian behavior, but also analyze factors that influence driver decisions. The primary objective of this project is to establish a framework for investigating pedestrian-motorist nonverbal "negotiations" at semi-controlled locations. This research focuses on the crosswalks on North University Street at Second Street (Figure 1(b) and (c)) on the Purdue University campus. This location was chosen because (1) it has a variety of crossing conditions and (2) it was converted from a one-way street in 2017 to two-way operation in 2018. Having a video record of pedestrian-motorist interactions permitted a detailed examination of those interactions. Furthermore, the change from one-way traffic to two-way traffic provided a rare opportunity to study the behavior of a similar population of pedestrians and motorists at a location that underwent a significant change. In this study, four primary research questions were pursued:

1. Which factors can describe and explain the pedestrian-motorist interactions at semi-controlled crossing locations?
2. How will the pedestrian-motorist interaction change if a one-way street is changed to two-way operation?
3. Which characteristics will determine how long a pedestrian waits at a semi-controlled crossing location?
4. How can pedestrian-vehicle interaction and pedestrian waiting behavior at semi-controlled crossing locations be modeled?

## 2. REVIEW OF LITERATURE

### 2.1. Pedestrian Crossing Behavior

Gap acceptance theory has been commonly used to model pedestrian decision-making. Probability-based approaches and modeling approaches are two main forms of pedestrian gap acceptance studies. Sun et al. (2002) proposed a Pedestrian Gap Acceptance method to model pedestrian decision strategies. They considered the probability of accepting a gap as a random variable that was obtained by fitting distributions to field data. Zhuang and Wu (2011) used statistical methods to analyze pedestrian crossing patterns (eye contact and running) and used gap acceptance models to describe pedestrian crossing behaviors at unsignalized crosswalks in China. Yannis et al. (2013) also employed a probability-based approach (lognormal regression) to test pedestrian gap acceptance in front of approaching vehicles at mid-block crossings. A binary logit model was used to explore the effect of gaps and other related parameters that affected pedestrian decision strategies. Kadali et al. (2014) compared the effectiveness of a non-linear model (artificial neural network) with a linear model (multivariate regression) in establishing a relationship between pedestrian gap acceptance behavior and explanatory factors.

These studies listed above were primarily based on pedestrian gap acceptance behavior. However, gap acceptance theory has some limitations and may be inadequate to explain pedestrian crossing behavior. Some researchers have explored the family of discrete choice models to describe pedestrian crossing strategies. Himanen and Kumala (1988) developed a multinomial logit model to interpret the "negotiations" between drivers and pedestrians on crosswalks. Papadimitriou et al. $(2012,2016)$ designed surveys to investigate the impact of human factors on pedestrian crossing decisions by means of principal component analysis. Furthermore, a theoretical framework to model pedestrian crossing decision making process in urban trips was proposed via different discrete choice models. Lord et al. (2018) designed a questionnaire to understand the relationships between crossing strategies and the perceptions of the elderly, and logistic regression models were applied to explain the observed behaviors.

### 2.2. Driver Yielding Behavior

Discrete choice models have been widely used to analyze driver behavior, considering several explanatory parameters under different traffic or concurrent conditions at unsignalized crosswalks. Schroeder and Rouphail $(2010,2013)$ used logistic regression to predict driver yielding behavior at "semi-controlled" crosswalks and roundabouts based on vehicle dynamics, pedestrian characteristics, and environmental conditions. Sucha et al. (2017) studied the communications between pedestrians and drivers with respect to physical gestures (waving and eye contact), then took advantage of logistic regression to explore factors that had an influence on driver yielding behavior. Cloutier et al. (2018) applied a mixed-effects logit model to evaluate factors related to the likelihood of interactions between pedestrians and motorists. Other researchers explored game theory to explain the interactions between pedestrians and vehicles (Guan et al., 2016; Bjørnskau, 2017; Camara et al., 2018).

### 2.3. Pedestrian Wait Behavior

Recently, researchers also considered pedestrian wait time as one of the most important performance metrics in pedestrian-motorist interaction. Survival models have been used in transportation studies, especially for travel time and wait time, because of their flexibility in dealing with duration-based data (Washington et al., 2010). Nonparametric, semi-parametric and fully parametric survival models have been utilized to explore the effects of human factors on pedestrian waiting behavior. The Kaplan-Meier estimator (Kaplan and Meier, 1958) and the Lee-Carter method (Lee and Carter, 1992) are two prevailing approaches for non-parametric survival models, which provide practical estimates of survival probabilities and a raw graphical representation of the survival distribution (Washington et al., 2010).

In transportation studies, the nonparametric Kaplan-Meier estimator was widely applied to investigate pedestrian wait duration before making unsafe crossings at signalized intersections (Tiwari et al. 2007; Guo et al, 2011). Cox (1972) developed a semi-parametric survival model duration model that included the effects of covariates. Guo et al. (2011) applied the semi-parametric Cox proportional hazard model to analyze influences of personal characteristics and external environment on pedestrian wait duration at signalized intersections in China based on both legal and illegal crossings. Instead of nonparametric and semi-parametric models, fully parametric models were developed by applying alternative statistical distributions for the baseline hazard function. Fully-parametric models are recently popular because their capacities of fitting different types of baseline hazard functions. Guo et al. (2012) further extended their previous research by applying both nonparametric and fully parametric models to explore the effects of human factors on pedestrian waiting behavior at signalized intersections. Li (2013) focused on pedestrian wait time at signalized intersections, and U-shaped distribution of pedestrian wait time was found. Yang et al. (2015) proposed hazard-based duration approach to study the wait time for cyclists and electronic bike riders.

There has been research on how long a pedestrian will wait at unsignalized crosswalks (Hamed, 2001). How long a pedestrian decides to wait reflects how safe he/she perceives it is to cross the roadway. Pedestrians may feel unsafe and tend to wait longer when motorists exhibit aggressive driving behaviors. An investigation of pedestrian crossing behaviors and wait durations at such locations can be useful in developing policies and control strategies to enhance a pedestrian's perceived safety, reduce pedestrian delay, and improve the level of service (LOS) of a unsignalized intersections.

Pedestrian waiting behavior consists of not only a single event, but also recurrent events. Survival models have limitations in dealing with recurrent events in a dataset. Traditional survival models consider only the first event as the critical event. However, the first event analysis only considers the first event, while the subsequent events are ignored. Secondly, the repeated events are treated as identical treatments, which should be modeled differently.

The Anderson-Gill (AG) model (Andersen and Gill, 1982) extended the Cox model to handle recurrent events by applying a counting process. They applied a baseline hazard function to all events. The subject of interest was the number of repeated events, given a specific period. The AG model is widely used in medical science, but it has a strong proportional odds assumption that, in practice, is difficult to be satisfied. Prentice, Williams, and Peterson (PWP) (1981) further recurrent events regression analysis by stratifying the events as ordered series, which allowed separate baseline hazard functions and coefficients to vary across events. Wei, Lin and Weissfeld (1989) also developed a model for a repeated events
modeling approach. However, this model is less efficient than PWP, due to its complicated nature. Although, AG, PWP and WLW are three classic models that have been widely investigated in the repeated events analysis, these three models focused only on the probability of occurrence, rather than the transition process between repeated events. Multi-State Models would be appropriate alternatives for recurrent events analysis.

There are three main research gaps in the existing research literature concerning pedestrianmotorist interaction. First, gap acceptance theory has been the prevailing method to analyze pedestrian behavior, but it may not be adequate at semi-controlled or controlled locations, where pedestrians can assert the priority to cross. Second, most research has studied either pedestrian behavior or motorist behavior separately. The research questions in our study call for an integrated framework that considers the potential relationships between pedestrian behavior and motorist behavior. Finally, the existing literature shows that survival models have the potential to model pedestrian waiting behavior at unsignalized semi-controlled crosswalks as a time-to-first-event analysis. Additionally, multi-state models can accommodate the dynamic modeling of pedestrian waiting behavior as recurrent event analysis. In the following sections, both time-to-first-event approach and recurrent events modeling approach are discussed.

## 3. RESEARCH METHODOLOGY

### 3.1. Study Site Description

## North University Street at Second Street

North University Street, at the T intersection with Second Street, has two lanes, each 12 ft wide, with a speed limit of $25 \mathrm{miles} / \mathrm{h}$. The two crosswalks are used by students and staff walking between central campus to the east and parking facilities and residences to the west. The first set of videos were made in Spring 2017, when North University Street was a one-way northbound street. By the time the second set of videos were made in Spring 2018, the streets had been converted to two-way operation. This conversion provided a rare opportunity to study pedestrian-motorist interaction at the same site under different conditions. The two sets of video recordings were made at four different time periods (7:40-8:20; 12:4013:25; 13:20-14:00; 16:20-17:00), when moderate traffic volumes and pedestrian flows were observed. The authors recorded approximately 3 hours of video for each set, resulting in a total of 3400 pedestrianmotorist interactions.

### 3.2. Definition

### 3.2.1. Interaction

The time-synchronized videos were processed in the laboratory. Interaction-based data were extracted to support the development of statistical models to investigate the "negotiation" between pedestrian and motorist. In this study, we define the interaction between pedestrian and motorist as the behavior of either party when in the area of influence of the other. The area of influence is defined by a vehicle being close enough to the crosswalk to affect the pedestrian's crossing decision. We assume (based on behavior seen in the video) that pedestrians make their crossing decisions within the curb area (within 2 meters of the street). From our observations, most pedestrians take a definite look for vehicles within the curb area, and most pedestrians wait within the curb area if drivers do not give an indication of yielding to them. Based on situations seen in the video recordings, an interaction can happen in several ways (Fricker and Zhang, 2019):

1. A pedestrian arrives at the curb and crosses immediately (without delay) while a vehicle accelerates, slows down or stops to avoid a conflict.
2. A pedestrian arrives at the curb and slows down or stops, but a vehicle slows down or stops to yield to the pedestrian.
3. A pedestrian arrives at the curb and slows down or stops, while a vehicle slows down, but does not yield to the pedestrian.
4. A pedestrian arrives at the curb and slows down or stops, while a vehicle keeps a constant speed or accelerates, not yielding to pedestrian.
An interaction does not occur if:
5. A pedestrian arrives at the curb area, but there is no vehicle close enough to the crosswalk to affect the pedestrian's crossing decision.
6. A vehicle approaches the crosswalk, but there is no pedestrian present who is attempting to cross.

### 3.2.2. First Event \& Critical Event

Based on the definition of interaction, we aimed to investigate pedestrian wait time when pedestrianmotorist interactions happen. A pedestrian can interact with either one vehicle or multiple vehicles, so that the pedestrian wait time dataset is mixed with single interaction events and recurrent interaction events. In survival models, the first event was considered as a "critical event", because it had greater impact on the pedestrian's crossing decision than the other vehicle(s) did. To investigate the impacts of critical vehicles on pedestrian behavior, we coded information for the critical vehicle: vehicle type, driving in the near or far lane, distance to pedestrian, and approach speed at the time when pedestrian reaches the curb.

### 3.2.3. Recurrent Events

Multi-state Semi-Markov models allow the estimation of the instantaneous impact of factors on the probability of transition from different states. By applying multi-state Markov model, we modeled transitions in Figure 2 in terms of three states:

1. A pedestrian reaches the curb area as a vehicle approaches, so that an interaction occurs.
2. Pedestrian rejects the lag (and, if necessary, subsequent gaps).
3. Pedestrian accepts the lag (or gap) and crosses the street.

The potential transition for this model is defined as:

- Transition 1-3: accept the lag directly (vehicle yields).
- Transition 1-2: reject the lag and await recurrent gaps (vehicle fails to yield).
- Transition 2-2: reject following gaps. (This transition is not considered in this thesis)
- Transition 2-3: accept a subsequent gap.

A pedestrian who rejects a lag (1-2) would not be considered for Transition 1-3. In contrast, an individual whose first transition was 1-3, is considered simultaneously with the transition as 1-2.

Transition $2 \rightarrow 2$


Figure 2.Multi-State Framework for Pedestrian Waiting Behavior (Zhang, 2019)

## 4. PEDESTRIAN-VEHICLE INTERACTION

### 4.1. Explanatory Variables

In order to explore pedestrian-vehicle interaction, all explanatory variables, including pedestrian characteristics and dynamics, vehicle dynamics, and environmental conditions are documented in Table 1. Whenever a pedestrian entered the curb area while a motorist was present in the area of influence, values for all the variables listed in Table 1 were manually recorded. For Hesitation parameter, the $75 \%$ percentile of pedestrian wait time for non-hesitation behavior is 1.57 s (one-way) while the $25 \%$ percentile pedestrian wait time for hesitation behavior is 1.735 s . The number for two-way case is 1.80 s and 2.51 s separately. We re-examine the $25 \%$ for both Hesitation and Non-hesitation behavior time after time through watching the videos. For CloseFollow parameter, we define that the object vehicle has close follower (CloseFollow $=1$ ), if the vehicle has a follower at a short headway of approximately 2-4 seconds, which has been defined similarly in former research literature (Schroeder and Rouphail, 2011).

### 4.1.1. Pedestrian Behavior

Based on the recorded interactions between pedestrians and motorists, a predictive model of pedestrian crossing behavior might be developed. There are two potential outcomes that describe pedestrian behavior:

- Pedestrian Crosses ( $\mathrm{Y}=1$ ): the motorist in the interacted vehicle provides an opportunity for the pedestrian to cross.
- Pedestrian Yields $(\mathrm{Y}=0)$ : a pedestrian offers the motorist an opportunity to pass through the crosswalk first in an interaction.
The variables that seemed appropriate to use in a model of pedestrian behavior are indicated by check marks in the "Pedestrian Model" column of Table 1.


### 4.1.2. Motorist Behavior

Based on numerous recorded interactions, a driver's likelihood to decelerate was determined to be a key factor in the negotiation between pedestrian and motorist. A driver's likelihood to decelerate had better explanatory power than likelihood to yield, because, in an interaction, a motorist could slow down initially, but have the pedestrian wave to the motorist to go first. In this situation, the driver's action to decelerate is considered an important element of the interaction, even if the motorist did not eventually yield to the pedestrian. Consequently, by assigning levels of deceleration to each motorist, the potential outcomes for a motorist in an interaction are:

- Level 1. Keep a constant speed or accelerate: a motorist does not slow down, and the interaction does not cause delays for the motorist.
- Level 2. Decelerate but do not fully stop: a motorist decelerates during an interaction but does not fully stop and incurs some delay.
- Level 3. A motorist stops to accommodate a pedestrian and incurs a delay that is usually greater than in Level 2.

Table 1.Explanatory Variables

| Parameters | Variable Description | Value | Pedestrian <br> Model | Motorist <br> Model |
| :---: | :---: | :---: | :---: | :---: |
| Pedestrian Characteristic and Dynamics |  |  |  |  |
| GroupSize | The number of pedestrians in the curb area, including the subject pedestrian | Integer | $\checkmark$ |  |
| AgeRange | Estimated age range for subject pedestrian(s) (1: 0-10; 2: 10-30; 3 : $30-50$; and 4: >50). | Indicators | $\checkmark$ |  |
| Sex | Sex of subject pedestrian | $\begin{aligned} & \hline \text { Male }=1 ; \\ & \text { Female }=0 \end{aligned}$ | $\checkmark$ | $\checkmark$ |
| Hesitation | Does the pedestrian slow down or wait at curb? | $\mathrm{Y}=1 / \mathrm{N}=0$ | $\checkmark$ | $\checkmark$ |
| Distraction | Does a pedestrian approach and/or cross while using a cellphone or talking? | $\mathrm{Y}=1 / \mathrm{N}=0$ | $\checkmark$ | $\checkmark$ |
| FlowWith | The number of pedestrians already crossing in the crosswalk in the same direction when subject pedestrian arrives at curb area | Integer | $\checkmark$ |  |
| FlowAgainst | The number of pedestrians already crossing in the crosswalk in the opposite direction when subject pedestrian arrives at curb area | Integer | $\checkmark$ |  |
| FlowOn | Total number of pedestrians already crossing in the crosswalk when an interaction occurs (FlowWith + FlowAgainst). | Integer |  | $\checkmark$ |
| SameDirec | The number of pedestrians present in the curb area crossing in the same direction as the subject pedestrian | Integer | $\checkmark$ |  |
| DiffDirec | The number of pedestrians present in a curb area with crossing direction opposite of the subject pedestrian | Integer | $\checkmark$ |  |
| PedWait | Total number of pedestrians waiting in the curb areas when an interaction occurs (SameDirec + DiffDirec) | Integer |  | $\checkmark$ |
| Vehicle Dynamics |  |  |  |  |
| ApprSpeed | The approach speed of interacted vehicles when a pedestrian enters the curb area. (mph) | Float | $\checkmark$ | $\checkmark$ |
| SlowsDown | Does a vehicle slow down or stop on the approach to the crosswalk when a pedestrian enters the curb area? | $\mathrm{Y}=1 / \mathrm{N}=0$ | $\checkmark$ |  |
| CloseFollow | Does the interacted vehicle have a close follower when an interaction occurs? | $\mathrm{Y}=1 / \mathrm{N}=0$ | $\checkmark$ | $\checkmark$ |
| AdjVeh | Is a vehicle already present in the adjacent lane when a motorist begins to interact with a pedestrian? | $\mathrm{Y}=1 / \mathrm{N}=0$ | $\checkmark$ | $\checkmark$ |
| Environmental Characteristics |  |  |  |  |
| Distance | The distance of interacted vehicle(s) to subject pedestrians when interaction begins. (in ft.) | Float | $\checkmark$ | $\checkmark$ |
| NoF | Is pedestrian entering curb area on the near side or far side of the approaching vehicle's lane? | $\begin{aligned} & \hline \text { Near=0; } \\ & \text { Far=1 } \end{aligned}$ | $\checkmark$ | $\checkmark$ |
| Response Behavior |  |  |  |  |
| Pedestrian Outcomes | Cross: $\mathrm{Y}=1$; Wait/Yield: $\mathrm{Y}=0$ | Indicators | $\checkmark$ |  |
| Vehicle Response | Level of vehicle deceleration when pedestrians enter crosswalks ( 3 = stops; 2 = slows down; 1 = Does not slow down). | Indicators |  | $\checkmark$ |

### 4.2. Descriptive Statistics

We examined descriptive statistics for general trends in the data. In videos of the one-way University Street, there were 1,759 interactions involving 1,133 pedestrians and 498 motorists. It was observed that some pedestrians interacted with more than one motorist, and vice versa. Of the total interactions, 1,240 (70.5 percent) resulted in pedestrians crossing, while 519 ( 29.5 percent) of total interactions were of the "Pedestrian Yield to Motorist" type. In 993 out of 1,759 cases ( 56.5 percent), motorists chose to slow down or stop for pedestrians. When University Street was in two-way operation, the number of interactions was 1,574 (involving 933 pedestrians and 506 motorists). Of the total interactions, 1,061 (67.4 percent) had pedestrians crossing, while 513 ( 32.3 percent) of total interactions were of the "Pedestrian Yield to Motorist" type. Moreover, in 1,005 out of 1,574 cases ( 63.9 percent), motorists chose to slow down or stop for pedestrians during interactions on the two-way street.

In Table 2, descriptive statistics of all explanatory variables are shown. The asterisks in the "Mean" columns indicate the level of significance of the difference between the mean value of variables for oneway operation and two-way operation, found using $t$ tests. Overall, the data showed a significantly higher percent of pedestrians hesitating (Hesitation, $54.9 \%$ ) on the one-way street than with two-way traffic $(49.9 \%)$. In one-way cases, $29.7 \%$ of vehicles have a vehicle following closely behind (CloseFollow) when they are involved in an interaction. However, the value of CloseFollow for two-way streets is $50.6 \%$. Also, in one-way cases, $43.4 \%$ of vehicles arrived at the study area with a vehicle present in the adjacent lane. However, this number for the two-way street is $52.1 \%$. The average value of the distance from vehicle to crosswalk is 74.6 ft . on the one-way street, which is significantly different from the two-way case ( 64.7 ft ).

Table 2.Descriptive Statistics

| Variables | One-way |  |  | Two-way |
| :--- | :--- | :--- | :--- | :--- |
|  | Mean | Std.Dev. | Mean | Std.Dev. |
| GroupSize | $2.53^{* * *}$ | 2.331 | $2.054^{* * *}$ | 1.483 |
| AgeRange | $2.193^{* * *}$ | 0.425 | $2.321^{* * *}$ | 0.542 |
| Sex | $0.429^{*}$ | 0.495 | $0.471^{*}$ | 0.499 |
| Hesitation | $0.549^{* *}$ | 0.498 | $0.499^{* *}$ | 0.5 |
| Distraction | 0.146 | 0.353 | 0.163 | 0.37 |
| FlowWith | 1.229 | 2.049 | 1.151 | 1.847 |
| FlowAgainst | 0.875 | 1.581 | 0.792 | 1.607 |
| FlowOn | 2.103 | 3.015 | 1.943 | 2.8 |
| SameDirec | 1.017 | 1.402 | 0.931 | 1.204 |
| DiffDirec | $0.629^{* *}$ | 1.123 | $0.525^{* *}$ | 1.003 |
| PedWait | 1.646 | 2.005 | 1.457 | 1.725 |
| AppSpeed | 8.543 | 6.939 | 8.355 | 7.794 |
| SlowsDown | $0.565^{* * *}$ | 0.496 | $0.639^{* * *}$ | 0.481 |
| CloseFollow | $0.297^{* * *}$ | 0.457 | $0.506^{* * *}$ | 0.5 |
| AdjVeh | $0.434^{* * *}$ | 0.496 | $0.521^{* * *}$ | 0.5 |
| Distance | $74.592^{* * *}$ | 54.729 | $64.673^{* * *}$ | 49.12 |
| NoF | 1.505 | 0.5 | 1.502 | 0.5 |
| Pedestrian Outcomes | 0.705 | 0.456 | 0.674 | 0.469 |
| Vehicle Response | $1.851^{* *}$ | 0.837 | $1.945^{* *}$ | 0.812 |
| $* \mathrm{p}<.05 ; * *$ p<.01; $* * * \mathrm{p}<.001$ |  |  |  |  |

### 4.3. Modeling Approach

### 4.3.1. Pedestrian Model

A pedestrian's Cross or Yield behavior has a binary outcome: Cross ( $\mathrm{Y}=1$ ) or Yield ( $\mathrm{Y}=0$ ). Commonly, a binary logistic regression model is applied to estimate the probability that a particular choice happened, based on a series of explanatory variables. Using this method, a linear model was built with explanatory variables by transforming the outcomes into $\operatorname{Prob}\{\mathrm{Y}=1\}$. The logistic regression model assumes that, for every explanatory property (Harrell, 2015),

$$
\begin{equation*}
\operatorname{logit}(Y=1 \mid X)=\log \left[\frac{1-P(Y=1)}{P(Y=1)}\right]=\sum_{i=1}^{I} \beta_{i} X_{i}+C \tag{1}
\end{equation*}
$$

where C is the intercept and $\beta_{i}$ is the change in the $\log$ odds per unit change in $X_{i}$, while all other variables are unchanged. Equation (2) can be used to describe the correlates between odds (Y) and variables (Harrell, 2015),

$$
\begin{equation*}
\operatorname{odds}(Y=1 \mid X)=\exp (X \beta) \tag{2}
\end{equation*}
$$

The regression parameters can also be written in terms of odds ratios. The odds that $Y=1$ when $X_{j}$ is increased by d, divided by the odds at $X_{j}$ is
odds ratio $=\frac{\operatorname{odds}\left\{Y=1 \mid X_{1}, X_{2}, \ldots, X_{j}+d, \ldots, X_{k}\right\}}{\operatorname{odds}\left\{Y=1 \mid X_{1}, X_{2}, \ldots, X_{j}, \ldots, X_{k}\right\}}=\exp \left[\beta_{j} X_{j}+\beta_{j} d-\beta_{j} X_{j}\right]=\exp \left(\beta_{j} d\right)$
The mixed effects logit model has been widely used in transportation safety research due to its flexibility in model structure. Compared with binary logistic regression, the mixed-effects logit model considers the probability as the integral of the standard logit model over a density distribution of a parameter (Ye et al., 2014). The mixed effects logit model can be written as:

$$
\begin{equation*}
P(Y=1)=\int \frac{\exp \left(\alpha_{k}+\beta_{k} X\right)}{\sum_{\forall k} \exp \left(\alpha_{k}+\beta_{k} X\right)} f(\beta \mid \theta) d \beta \tag{4}
\end{equation*}
$$

The $\theta$ S in the model are normally distributed in both one-way case and two-way case. Estimated values are shown in Table 3. Normally, the mixed effects logit model is compared with binary logistic regression together and AIC is a critical indicator for model selection, which balance the fitness and model complexity. The AIC can be expressed as (Akaike, 1987):

$$
\begin{equation*}
A I C=-2 \ln (\text { likelihood })+2 k \tag{5}
\end{equation*}
$$

where $k$ is the number of parameters. In this part of the study, both binary logistic regression and mixed effects logit model were tried. The model that best represented the data was chosen based on the AIC. The model results are shown in Table 3.

Table 3 Binary Logistic Regression Results for Pedestrian Models

| Variables | One-way |  | Two-way |  | Interacted Coefficients of Combined Data |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Logistic Model | MixedEffects Logit | Logistic Model | MixedEffects Logit | Logistic Model | MixedEffects Logit |
| GroupSize | - | - | 0.443** | 0.452** | 0.424* | 0.424* |
| AgeRange | - | - | - | - | - | - |
| Sex | - | - | - | - | - | - |
| Hesitation | -4.767*** | -7.878*** | $-3.699^{* * *}$ | -3.788*** | - | - |
| Distraction | - | - | - | - | - | - |
| FlowWith | 0.331*** | 0.522** | - | - | -0.456*** | -0.573*** |
| FlowAgainst | 0.28** | 0.425* | 0.227* | 0.232* | - | - |
| SameDirec | - | - | - | - | - | - |
| DiffDirec | - | - | 0.384** | 0.391** | - | - |
| AppSpeed | -0.147*** | $-0.224^{* * *}$ | -0.111*** | -0.114*** | - | - |
| SlowsDown | 2.593*** | 4.352*** | 3.03*** | 3.091*** | - | - |
| CloseFollow | - | - | - | - | - | - |
| AdjVeh | -0.546** | -1.033** | - | - | 0.634* | 0.865* |
| Distance | 0.042*** | 0.069*** | 0.036*** | 0.036*** | - | - |
| NoF | -0.689*** | -0.932** | -0.65** | -0.665** | - | - |
| Constant | 4.604*** | 6.647*** | - | - | 3.209*** | 3.972*** |
| Log <br> Likelihood | -336.817 | -320.589 | -328.075 | -328.008 | -658.987 | -648.043 |
| $\theta$ | - | Esti: 2.750 <br> std: 0.508 | - | $\begin{aligned} & \text { Esti: } 0.0025 \\ & \text { std: } 3.806 \end{aligned}$ | - | Esti: 1.47 <br> std: 0.253 |
| AIC | 693.6333 | 663.1773 | 676.1501 | 678.0157 | 1383.973 | 1364.087 |
| BIC | 748.3583 | 723.3748 | 729.7638 | 736.9908 | 1585.657 | 1571.882 |
| Pseudo R ${ }^{2}$ | 0.6843 | - | 0.6698 | - | 0.6805 | - |
| Observations | 1759 |  | 1574 |  | 3333 |  |
| * p < 0.05; ** p < 0.01; *** p < 0.001 |  |  |  |  |  |  |

### 4.3.2. Pedestrian Model Discussion

From Table 3, it can be seen that significant variables were estimated similarly in both models. For oneway data, mixed logit works slightly better based on the AIC and BIC. However, for two-way operation, binary logistic regression exhibits superior performance. Due to the similarity of results from two models, to avoid confusion, we mainly discuss results and findings based on the binary logistic regression model.

The significant variables in the logistic regression model (Table 3) suggest that, in a pedestrian-motorist interaction, pedestrians are more likely to decide to cross under the following conditions:
a. If a pedestrian is assertive without hesitation (Hesitation). Pedestrian-motorist interactions were compared with only the Hesitation variable changing, while keeping other variables equal to their average values. If a pedestrian slows down at the curb while interacting with a motorist, the coefficient indicates that the probability of pedestrian crossing with Hesitation $=1$ is,

$$
\begin{align*}
\sum_{i=1}^{I} \beta_{i} X_{i, 1}+C= & -4.767 * 1+0.331 * 1.229+0.28 * 0.875-0.796 * 2.19-0.147 * 8.543+2.593 * 0.566  \tag{6}\\
& -0.546 * 0.434+0.042 * 74.592-0.689 * 1.505+4.604=0.81633
\end{align*}
$$

$$
\begin{equation*}
P(Y=1 \mid \text { Hesitation }=1)=\frac{e^{\sum_{i=1}^{I} \beta_{i} x_{i, 1}+C}}{1+e^{\sum_{i=1}^{I} \beta_{i} X_{i, 1}+C}}=0.693=69.3 \% \tag{7}
\end{equation*}
$$

The probability of a pedestrian crossing under the same conditions, but with Hesitation $=0$, is

$$
\begin{align*}
\sum_{i=1}^{I} \beta_{i} X_{i, 2}+C= & -4.767 * 0+0.331 * 1.229+0.28 * 0.875-0.796 * 2.19-0.147 * 8.543+2.593 * 0.566  \tag{8}\\
& -0.546 * 0.434+0.042 * 74.592-0.689 * 1.505+4.604=5.5833 \\
P(Y=1 \mid \text { Hesitation }=0)= & \frac{e^{\sum_{i=1}^{I} \beta_{i} X_{i, 2}+C}}{\sum_{\sum_{i=1}^{I} \beta_{i} X_{i, 2}+C}}=0.996=99.6 \% \tag{9}
\end{align*}
$$

b. If a driver decelerates (SlowsDown). If a driver slows down during an interaction, the probability of a pedestrian crossing is $98.4 \%$, while the probability for a non-slowing down event is $81.9 \%$. This is a clear indication of a motorist yielding to a pedestrian during the interaction. The effects of the Hesitation and SlowsDown variables are shown in Figure 3 in terms of Distance.


Figure 3.Effects of Hesitation and Slow Down on Pedestrian Crossing Behavior
c. If a car is approaching at a lower speed (AppSpeed). If the approach speed of a vehicle is 10 mph , the probability of pedestrian crossing is $71 \%$, while the probabilities are $36.0 \%$ for 20 mph and $11.4 \%$ for 30 mph . Other studies (e.g., Brüde and Jörgen, 1993; Leaf and Preusser, 2006) found similar relationships.
d. If the distance (in ft ) between a pedestrian and motorist is great (Distance). If the interaction distance between vehicle and object crosswalk is 20 ft ., the probability of pedestrian crossing is $66.5 \%$, while the probabilities are $91.4 \%$ for 60 ft . and $98.3 \%$ for 100 ft . The effects of Distance and AppSpeed variables can be seen in Figure 4.


Figure 4.Effects of Distance and Speed on Pedestrian Crossing Behavior
e. If there is no vehicle in the adjacent lane (AdjVeh). If pedestrian has to interact with two vehicles in different lanes, the probability of a pedestrian crossing is $93.5 \%$, compared with the case in which there is no adjacent vehicle (object pedestrian interacts with only one vehicle in any lane) $(96.1 \%)$. Schroeder and Rouphail (2011) found a similar relationship.
f. If a vehicle is in the near lane ( $\mathrm{NoF}=0$ ). The probability of a pedestrian deciding to cross is $96.5 \%$ when a far lane interaction occurs, to $98.2 \%$ with a near-lane interaction. A plausible explanation is that, all else being equal, a pedestrian is more confident when crossing before a vehicle in the near lane arrives, compared with the longer crossing distance to the far lane and the risk of being trapped in the crosswalk while waiting for a far lane vehicle to yield or proceed. Effects of AdjVeh and NoF variables can be seen in Figure 5 in terms of Distance.


Figure 5.Effects of AdjVeh and NoF on Pedestrian Crossing Behavior
g. If other pedestrians are using the crosswalk (FlowWith and FlowAgainst) (Figure 6).


Figure 6.Effects of FlowWith and FlowAgainst on Pedestrian Crossing Behavior
h. If other pedestrians are using the crosswalk (FlowWith and FlowAgainst). The effects are shown in Figure 6.

Some of the findings in this study are similar to individual findings found in other studies; other findings in this study represent new contributions. All of the findings are plausible. This gives the model credibility, subject to a more careful look at the model's values, after examining related models.

Based on the data for two-way University Street, the binary logistic regression model (Table 3) suggests that, in a pedestrian-motorist interaction, pedestrians are more likely to cross ...
i. if pedestrian acts without hesitation (Hesitation=0).
j. if a driver decelerates (SlowsDown=1). See Figure 7.
k. if a vehicle approaches at a lower speed (ApprSpeed).

1. if the distance between a pedestrian and motorist is greater (Distance). See Figure 8.
m . if the other pedestrians are crossing in the same direction as the pedestrian being observed (FlowWith).
n. if an interacted vehicle is in the near lane ( $\mathrm{NoF}=0$ ).
o. if there is a pedestrian waiting on the opposite side of the street (DiffDirec) See Figure 9.
p. if a pedestrian is grouped with other people (GroupSize) See Figure 9.


Figure 7.Effects of Hesitation and SlowsDown on Pedestrian Crossing Behavior


Figure 8.Effects of Distance and Speed on Pedestrian Crossing Behavior


Figure 9.Effects of DiffDirec and GroupSize Variables on Pedestrian Crossing Behavior

After University Street was converted to two-way operation, fewer parameters were found to be statistically significant. The first four variables listed above for two-way operation were also significant for the one-way data. This could be interpreted in the same way as in one-way case.

In an attempt to analyze the effect of the change on University Street from one-way to two-way, the data were combined and a group dummy variable (one-way case $=0$; two-way case $=1$ ) was introduced. The binary logit regression was run on the combined data, using an interaction term that pairs group dummy variables with independent variables, as shown in Equation 7 (Cross Validate, 2018):
$\operatorname{logit}(Y=1 \mid X)=\beta_{0}+\beta_{1} *$ indepVar $+\beta_{2} *$ groupDummy $*$ indepVar $+\beta_{3} *$ groupDummy
Where:

- $\beta_{1}$ is the vector of coefficients for the one-way case.
- $\beta_{2}$ is the vector that measures the difference in the coefficients between the two separate models (one-way and two-way).
- $\quad \beta_{3}$ shows the differences in intercepts between the separate models.

The second rightmost column in Table 3 shows the results for the coefficients of interest in the binary logistic regression model and represents elements of $\beta_{2}$ in Equation 7. Consequently, one can test whether each element in $\beta_{2}$ is significant, to show if the change from one-way to two-way has caused a significant change in the pedestrian crossing models. An example is that the DiffDirec estimated coefficients in one-way case is 0.133 with $95 \%$ confidence interval [ $-0.044450,0.311147$ ], while the number is $0.384^{* *}$ with $95 \%$ confidence interval [ $0.1317962,0.6501222$ ]. The $95 \%$ confidence interval for the estimated coefficients are overlapped so that in the second right most column in Table 3 is not reported as statistical significance. Other results show that:

- The coefficient for pedestrian arriving group size (GroupSize) has changed significantly. In the one-way case, it was negative and not statistically significant; for two-way case, it is positive and statistically significant ( $0.424^{* * *}$ in Table 3 ).
- There is a significant change in the effect of the FlowWith factor ( $-0.573 * * *$ in Table 3). In the one-way case, if there are already pedestrians in the crosswalk, pedestrians are more likely to cross. However, on a two-way street, the FlowWith factor had no significant impact on pedestrian crossing behavior.
- The impact of the presence of an adjacent vehicle is significant in the one-way case, but disappears in the two-way case, because the coefficient changes significantly ( $-0.685^{* *}$ in the table).


### 4.3.3. Motorist Model

By state law, at a semi-controlled crossing, a motorist is supposed to yield to a pedestrian who is "within the crosswalk". In both the one-way and two-way Pedestrian Model, a major factor in a pedestrian's decision to cross was the deceleration of the vehicle(s) during an interaction. Whether a vehicle slows down is a vital factor to study in the negotiations between pedestrians and motorists. In this section, we focus on the parameters that may have significant impacts on drivers' slowing down behavior. By assigning levels of deceleration intensity to each motorist (Level 1 - Motorist does not slow down; Level 2 -- Motorist slows down but does not stop; and Level 3 -- Motorist stops), an ordered logistic regression can be used to analyze the probability of a particular response level for a series of given parameters (Williams, 2016):
$\operatorname{Prob}\left(Y_{i}>j\right)=O(X \beta)=\frac{\exp \left(\alpha_{j}+X_{i} \beta_{j}\right)}{1+\exp \left(\alpha_{j}+X_{i} \beta_{j}\right)}, \quad j=1,2, \ldots M-1$
where $Y_{i}$ is the response variable (motorist action in this application), $M$ denotes the number of ordinal dependent variables ( $M=3$ levels here), and $\beta_{j}$ are the same for all categories, but $\alpha_{j}$ are not necessarily the same among categories. A critical assumption associated with the ordered logit model is the proportional odds assumption, which imposes the restriction that regression parameters (except constants) are the same across different dependent levels. However, for deceleration intensity, it is not clear whether distances between adjacent deceleration levels are equal. Considering that the proportional odds assumption may be violated by only a subset of variables, a generalized ordered logistic regression (GOLR) partial proportional odds model was adopted. Compared with the ordered logistic regression model, the GOLR model relaxes the proportional odds assumptions for some explanatory variables, while maintaining them for the variables that satisfied the proportional odds assumption (Williams, 2016). The model could be further revised using Equation 9 as:
$\operatorname{Prob}\left(Y_{i}>j\right)=g(X \beta)=\frac{\exp \left(\alpha_{j}+X_{1 i} \beta_{1}+X_{2 i} \beta_{2 j}\right)}{1+\exp \left(\alpha_{j}+X_{1 i} \beta_{1}+X_{2 i} \beta_{2 j}\right)}, \quad j=1,2, \ldots M-1$
where $\beta_{I}$ is the vector of variables that are subject to the proportional odds assumption. Explanatory variables $X_{2 i}$ that do not satisfy this assumption need the addition of coefficients $\beta_{21}$ to relax the proportional odds assumption. The results are shown in Table 4.

Table 4.Generalized Ordered Logistic Regression Model Results

| Variables | One-way |  | Two-way |  | Tests of equality of Coefficients (p-value) |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Coefficient Between |  | Coefficient Between |  |  |  |
|  | 1 and 2 | 2 and 3 | 1 and 2 | 2 and 3 | 1 and 2 | 2 and 3 |
| - | - | - | - | - | - | - |
| Sex | - | - | - | - | - | - |
| Distance | - | - | 0.007*** | - | 0.0035** | - |
| FlowOn | 0.137*** | 0.137*** | 0.101** | 0.101** | - | - |
| PedWait | - | - | - | - | - | - |
| AppSpeed | -0.304*** | -1.248*** | -0.302*** | -0.933*** | - | - |
| CloseFollow | -0.355* | -0.355* | - | - | - | - |
| AdjVeh | -0.569*** | -0.569*** | - | - | 0.0001*** | 0.0001*** |
| Hesitation | -0.384* | -0.384* | $-1.518^{* * *}$ | - | 0.0000*** | 0.0098** |
| Distraction | - | - | 0.451* | 0.451* | - | - |
| NoF | - | - | -0.3* | -0.3* | 0.0093** | - |
| Cutoffs | 4.44*** | 4.046*** | 4.114*** | 2.232*** | - | - |
| Log Likelihood | -784.588 |  | -754.676 |  | - | - |
| Pseudo R ${ }^{2}$ | 0.5855 |  | 0.5626 |  | - |  |
| Observations | 1759 |  | 1574 |  | - |  |
| * $\mathrm{p}<0.05$; ** $\mathrm{p}<0.01$; *** $\mathrm{p}<0.001$; "- = not applicable" |  |  |  |  |  |  |

In Table 4, the first column of coefficients can be interpreted in terms of Equation 9, where the dependent variable is recoded as $\operatorname{Prob}\left(\mathrm{Y}_{\mathrm{i}}>1\right)$, which is equivalent to the probability that motorist deceleration Levels 2 and 3 occur, i.e., $\mathrm{j}>1$. The second column of coefficients can be interpreted in terms of Equation 9 , where the dependent variable is recoded as $\operatorname{Prob}\left(\mathrm{Y}_{\mathrm{i}}>2\right)$, which is equivalent to the probability that motorist deceleration Level 3 occurs (Williams, 2016). This model has been widely used in traffic crash analysis to analyze the relationship between the severity of injury and associated variables. Furthermore, in some literature (e.g., Wang et al., 2008; Michalaki et al., 2015), marginal effects were used to measure the effect that a change in an explanatory variable has on the predicted probability of a specific category. The marginal effects are shown in Table 5.

Table 5.Marginal Effects

| Variables | One-way |  |  | Two-way |  |  |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- |
|  | Level 1 | Level 2 | Level 3 | Level 1 | Level 2 | Level 3 |
|  |  |  |  |  |  |  |
| Sex | - | - | - | - | - | - |
| Distance | - | - | - | $-0.007^{* * *}$ | $0.001^{* * *}$ | - |
| FlowOn | $-0.0176^{* * *}$ | $0.0146^{* * *}$ | $0.003^{* * *}$ | $-0.009^{* *}$ | $0.0041^{* *}$ | $0.0058^{* *}$ |
| PedWait | - | - | - | - | - | - |
| AppSpeed | $0.0393^{* * *}$ | $-0.015^{* * *}$ | $-0.0278^{* * *}$ | $0.0296^{* * *}$ | $0.0237^{* * *}$ | $-0.0533^{* * *}$ |
| CloseFollow | $0.0459^{*}$ | $-0.038^{*}$ | $-0.0079^{*}$ | - | - | - |
| AdjVeh | $0.0735^{* * *}$ | $-0.0609^{* * *}$ | $-0.0127^{* * *}$ | - | - | - |
| Hesitation | $0.0496^{*}$ | $-0.0411^{*}$ | $-0.0085^{*}$ | $0.1488^{* * *}$ | $-0.1602^{* * *}$ | - |
| Distraction | - | - | - | $-0.0442^{*}$ | $0.0184^{*}$ | $0.0258^{*}$ |
| NoF | - | - | - | $0.0294^{*}$ | $-0.0123^{*}$ | $-0.0171^{*}$ |
|  |  |  |  |  |  |  |
| *p $<0.05 ; * * \mathrm{p}<0.01 ; * * * \mathrm{p}<0.001 ;$ "-= not applicable" |  |  |  |  |  |  |

### 4.3.4. Motorist Model Discussion

For University Street in its one-way operation, the generalized ordered logistic regression model suggests that, in a pedestrian-motorist interaction:

- A motorist is more likely to slow down if the driver's approach speed is lower. In the GOLR model, the coefficients for both Level 1 and Levels 2 and 3 are negative ( -0.304 and -1.248 ) and marginal effects also suggest that a higher approach speed leads to a higher likelihood of non-slowing down behavior ( $0.0393 * * *$ in Table 5).
- A driver is more likely to slow down if there are no other vehicles present in the adjacent lane (AdjVeh Coef. $=-0.569 ; p$-value $=0.000$ in Table 4). Furthermore, the marginal effect of this
factor for Not Slowing Down (Level 1) is $0.0735^{* * *}$ in Table 5, which is a significant increase. This means that, for the one-way case, the behavior of a vehicle in the adjacent lane will cause significant effects on a motorist's decision.
- A motorist is less likely to decelerate if there is a close follower behind him/her (CloseFollow Coef. $=-0.355^{*}$ in Table 4). The marginal effects reflect that a driver will be more aggressive without slowing down ( $0.0459 *$ in Table 5) and less likely to brake or stop ( $-0.038^{*}$ and $-0.0079 *$ in Table 5) if another driver closely follows him/her.
- If a pedestrian slows down or stops at the curb during an interaction (Hesitation=1, Coef. $=-$ $0.384^{* * *}$ ), the marginal effects indicate that a driver will be more likely to continue without slowing down ( $+0.0496 *$ in Table 5) and less likely to slow down ( $-0.0411^{*}$ in Table 5).
- if an interaction occurs when there is a greater number of pedestrians already in the crosswalk (FlowOn), a driver is more likely to slow down (Coef. $=0.137 * * *$ in Table 4), because in Table 5, marginal effects indicate a positive impact on slowing down ( $0.0146^{* *}$ ) and on stopping ( $0.003^{* *}$ ), with negative impacts on non-deceleration $\left(-0.0176^{* * *}\right)$. It is intuitive that more pedestrians already in the crosswalk will lead to drivers slowing down.

For two-way University Street, the model shows the same effects as in the one-way case for variables FlowOn, AppSpeed, and Hesitation. Some other variables became significant.

- For the NoF variables in Table 4, their coefficients are $-0.3^{*}$, which means that the probability of a motorist slowing down is lower when the pedestrian is on the far curb ( $\mathrm{NoF}=1$ ), not the near curb, which is proved by the marginal effects (0.0294* for Level 1; -0.0123* for Level 2; and -0.0171* for Level 3 in Table 5).
- Distance $\left(\right.$ Coef. Distance $\left.=0.007^{*}\right)$ has significant impact on the driver's decision to slow down. With the increase of distance, the marginal effects for this parameter have a positive influence on a driver slowing down $\left(0.001^{* * *}\right)$ while having a negative effect on not-slowing down behavior (-0.0007***).
- For the distraction variable (Distraction), a driver is more likely to decelerate if a pedestrian in the interaction uses a cellphone or talks to others (Distraction, Coef. $=0.451 *$ in Table 4). In Table 5 , the marginal effects for this parameter have a positive influence on a driver slowing down and stopping ( $0.0184^{*}$ and $0.0258^{*}$ ), while having a negative effect on not-slowing down behavior (0.0442*).


### 4.3.5. Summary

As with the pedestrian model, we test whether the coefficients in the one-way and two-way cases are equal. The two rightmost columns in Table 4 show the p values for the hypothesis tests, which indicated that variables Hesitation, AdjVeh, NoF, and Distance change significantly. Meanwhile, compared with one-way street operation, some variables (CloseFollow and AdjVeh) were no longer significant. One interesting result is that, for one-way operation, driver behavior is influenced greatly by both pedestrian characteristics (Hesitation and FlowOn) and vehicle dynamics (AppSpeed, CloseFollow, and AdjVeh). For two-way operation, driver behavior is significantly determined by more pedestrian characteristics factors (Hesitation, Distraction, and FlowOn), fewer vehicle dynamics factors (AppSpeed) and more environmental characteristics factors (Distance and NoF), when an interaction occurs.

There exist limitations in the models. Endogeneity is introduced when we use SlowDown parameter in pedestrian model. Moreover, we explored the pedestrian and motorist behavior separately rather than study the complex game between pedestrian and motorist, which required a more complicated framework.

## 5. PEDESTRIAN WAIT TIME - SURVIVAL ANALYSIS

### 5.1. Accelerated Failure Time (AFT) Model

Based on the definition of interaction, we aimed to investigate pedestrian wait time when pedestrianmotorist interactions happen. A pedestrian can interact with either one vehicle or multiple vehicles, so that the pedestrian wait time dataset is mixed with single interaction events and recurrent interaction events. In survival models, the first event was considered as a "critical event", because it had greater impact on the pedestrian's crossing decision than the other vehicle(s) did. The object vehicle occurring in the first event is defined as the critical vehicle. To investigate the impacts of critical vehicles on pedestrian behavior, we coded information for the critical vehicle: vehicle type, driving in the near or far lane, distance to pedestrian, and approach speed at the time when pedestrian reaches the curb, as listed in Table 6.

Table 6 Explanatory Variables and Descriptive Statistics

| Variable | Description |
| :---: | :---: |
| Explanatory Variable |  |
| Pedestrian Wait Time | Duration (in seconds) between the time a pedestrian reached the curb area and the time the pedestrian started crossing (One-way: mean $=2.67 \mathrm{~s}, \mathrm{sd}=2.61 \mathrm{~s}$; Two-way: mean $=3.02 \mathrm{~s}, \mathrm{sd}=3.08 \mathrm{~s}$ ) |
| Independent Variable |  |
|  | Pedestrian characteristics |
| Sex | 1 if the pedestrian is Male (51.1\%), 0 if the pedestrian is Female (48.9\%). |
| Estimated Age Category | Young for pedestrians that appear to be younger than 30 years old (77\%); Mid-age for pedestrians between 30 and 50 years old ( $21 \%$ ); Elderly for pedestrians older than 50 years old ( $2 \%$ ). |
| Cellphone Indicator | 1 if pedestrian is using cellphone when waiting at the curb (9\%), 0 otherwise ( $91 \%$ ). |
| Talking Indicator | 1 if pedestrian is talking to others when waiting at the curb (8.4\%), 0 otherwise ( $91.6 \%$ ). |
| Traffic Condition |  |
| Vehicle Arrival Rate | Number of vehicles driving past the crosswalk per minute (mean $=8.72 \mathrm{veh} / \mathrm{min}, \mathrm{sd}=3.95 \mathrm{veh} / \mathrm{min}$ ) |
| Near Side Indicator | 1 if the critical vehicle is in the near lane (53\%), 0 if the critical vehicle is in the far lane (47\%). |
| Bus/Truck Indicator | 1 if the critical vehicle is bus or large truck ( $16.7 \%$ ), 0 otherwise (83.3\%). |
| Veh-to-Ped Distance | Distance (in ft ) between pedestrian and the first approaching vehicle (mean=61.8ft, $\mathrm{sd}=51.7 \mathrm{ft}$ ). |
| Approaching Speed | Speed (ft/s) of the first approaching vehicle when the pedestrian arrives at the curb (mean=12.4ft/s, $\mathrm{sd}=11.2 \mathrm{ft} / \mathrm{s}$ ). |
| Adjacent Vehicle Indicator | 1 if there is one or more vehicles presenting in the adjacent lane within the area of influence when the motorist begins to interact with a pedestrian ( $32.3 \%$ ), 0 otherwise ( $67.7 \%$ ). |
| Vehicle Close Follower Indicator | 1 if there is at least one vehicle closely following the critical vehicle (31.9\%), 0 otherwise (68.1\%). |
| Other Pedestrians |  |
| Group Size | Number of people in the pedestrian group ( $45.7 \%$ pedestrians came alone, $54.3 \%$ came in a group; mean $=2.31$, sd=2.03) |
| Nr Ped Waiting | Number of pedestrians waiting at the curb as a pedestrian arrives at the curb ( $52.8 \%$ cases of no pedestrians waiting, $47.2 \%$ cases of at least one pedestrian was waiting at the curb; mean $=0.84$, $\mathrm{sd}=1.21$ ) |
| Nr Ped Crossing | Number of pedestrians crossing the street as a pedestrian arrives at the curb ( $52.8 \%$ cases of no pedestrians crossing, $47.2 \%$ cases of at least one pedestrian crossing in crosswalk; mean=1.27, sd=1.99) |

Note: mean=average value; $s d=$ standard deviation

### 5.1.1. Log-Linear Model

First, a multivariate model was developed to analyze the relationship between pedestrian wait time and explanatory variables. At semi-controlled locations, pedestrian wait time is always non-negative. As a result, we transform the pedestrian wait time into log format. The log-linear model is:

$$
\begin{equation*}
\log \left(y_{i}\right)=\beta_{0}+\sum_{\forall k} \beta_{k} * x_{i, k}+\varepsilon_{i} \tag{10}
\end{equation*}
$$

where $\mathrm{y}_{\mathrm{i}}$ is the wait time and $x_{i, k}$ is the k -th variable for pedestrian $\mathrm{i} ; \beta_{0}$ is the estimated constant; $\beta_{k}$ is the coefficient estimated for the k -th variable, and $\varepsilon_{i}$ is the error term.

### 5.1.2. AFT Model Structure

Hazard-based duration models are widely used with duration-related datasets. Studies related to accident analysis (Nam and Mannering, 2000), travel activity behavior (Yang et al., 2015), and queueing theory usually applied hazard-based duration models due to their flexibility with time-dependent data.

Duration analysis primarily focuses on the length of time that elapsed from the starting state of an event until the ending state of an event. (In this study, we call the event an interaction.) Duration analysis is also interested in the likelihood that an event would end in the next short period of time, given its current state (Nam and Mannering, 2000). The hazard function at time (t) can be expressed as a density function $f(t)$, and its cumulative distribution function $F(t)$.

$$
\begin{equation*}
h(t)=\frac{f(t)}{1-F(t)}=\frac{f(t)}{S(t)}=\lim _{d t \rightarrow 0} \frac{\operatorname{Pr}(t \leq T \leq t+d t)}{d t * S(t)}=-\frac{S^{\prime}(t)}{S(t)} \tag{11}
\end{equation*}
$$

In this section, we discuss the fully parametric approach to investigate pedestrian waiting duration at semi-controlled locations. Different from non-parametric risk (hazard) functions, parametric functions need to be given a probability distribution $h(t)$. Typical probability distributions such as exponential, Weibull, log-logistic, log-normal, or Gompertz were investigated as alternatives in fully parametric hazardbased functions. The exponential distribution assumes that the hazard function is constant over time. Weibull or Gompertz distributions both assume that the hazard function is decreasing or increasing over time non-monotonically. The log-logistic distribution is a widely used probability distribution in hazardbased duration analysis, because of its flexibility in dealing with non-monotonic relationships. Figure 3 shows the matches between the empirical survival curves from the current data and the fitted curves using the candidate distributions. Log-logistic distribution and log-normal distribution were found to provide better fits among the candidate distributions.

These candidate distributions were tested in the AFT model, and their model performances were compared using the Akaike information criterion (AIC). The AIC is an estimator of the relative quality of statistical models, which provides a means of model selection. The AIC can be calculated using Equation (13), which represents a trade-off between model fit and model complexity (Akaine, 1987):

$$
\begin{equation*}
A I C=-2 * \log (\text { likelihood })+2(p+k) \tag{13}
\end{equation*}
$$

A lower AIC value indicates a truly better description of the data. According to the estimation results of the developed AFT models, the log-logistic model was found to provide better model performance (lower AIC) than other distributions in modeling pedestrian wait time. The hazard function $h(t)$, survival function $S(t)$, and survival (wait) time $T$ for a log-logistic AFT model are shown in Equations (14) to (17).


Figure 10.Non-Parametric Survival Estimations and Fitted Distributions
$h(t)=\frac{\lambda p t^{p-1}}{1+\lambda t^{p}}$
$T=\left(\frac{1}{S(t)}-1\right)^{\frac{1}{p}} \frac{1}{\lambda^{1 / p}}$
$T=\left(\frac{1}{S(t)}-1\right)^{\frac{1}{p}} \frac{1}{\lambda^{1 / p}}$
$T=\left(\frac{1}{S(t)}-1\right)^{\frac{1}{p}} \frac{1}{\lambda^{1 / p}}=\left(\frac{1}{S(t)}-1\right)^{\frac{1}{p}} \exp \left(\beta_{0}+\beta X\right)$
where $\frac{1}{\lambda^{1 / p}}$ is estimated by covariates considered in the model.

### 5.2. Model Discussions

The estimation results of the log-linear regression model and the survival models are presented in Table 7. The estimation results of the two models are mostly consistent with each other in terms of variable significance and sign consistency. The interpretation of the estimation results of the log-linear regression model is straightforward through the marginal effect: an influential variable with a positive sign indicates an increase on the wait time, meaning a one unit increase of an influential variable with coefficient $\beta_{i}$ leads to a $\beta_{i} * 100$ percent increase in pedestrian delay. For example, the coefficient estimated for Male is -0.0673 , meaning that the wait duration of a male pedestrian is $6.73 \%$ shorter than the duration of a female pedestrian, when all other factors are the same.

Table 7. Log-Linear Regression

| Log-Linear Model | One-way |  |  | Two-way |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Number of pedestrians | 1132 obs |  |  | 927 obs |  |  |
| Explanatory Variable | $\beta$ | t | P-value | $\beta$ | t | P-value |
| (Intercept) | -0.3294 | 0.149 | 0.027* | 0.69 | 0.1692 | <0.001*** |
| Sex | - | - | - | -0.0673 | 0.0383 | 0.079 . |
| Age | 0.1844 | 0.0413 | $<0.001^{* * *}$ | 0.0782 | 0.037 | 0.035* |
| Distraction | - | - | - | 0.1409 | 0.0529 | 0.008** |
| Group Size | -0.0228 | 0.0958 | 0.017* | -0.0495 | 0.0145 | 0.002** |
| Near or Far Side | - | - | - | -0.0974 | 0.0389 | 0.012* |
| Vehicle Type | 0.1487 | 0.0413 | $<0.001^{* * *}$ | - | - | - |
| Distance | 0.0017 | 0.0011 | 0.1256 | - | - | - |
| (Distance)^2 | -1.87E-05 | 5.76E-06 | 0.0012** | - | - | - |
| Approaching Speed | 0.0226 | 0.0055 | <0.001*** | 0.0106 | 0.0056 | 0.056. |
| (Approaching Speed)^2 | -0.00047 | 0.0002 | 0.079 . | - | - | - |
| Adjacent Vehicle | 0.2845 | 0.0416 | $<0.001^{* * *}$ | 0.2345 | 0.0427 | <0.001*** |
| Close Follower | 0.2751 | 0.0419 | $<0.001^{* * *}$ | 0.1895 | 0.0149 | <0.001*** |
| Nr Ped. Waiting | 0.0496 | 0.0113 | $<0.001^{* * *}$ | - | - | - |
| Nr Ped. Crossing | -0.0542 | 0.0071 | $<0.001^{* * *}$ | -0.0541 | 0.0085 | <0.001*** |
| $R^{\wedge} 2$ | 0.2418 |  |  | 0.1635 |  |  |
| $F$-statistic: | 25.45 |  |  | 12.73 |  |  |
| Degree of freedom | 11 |  |  | 10 |  |  |
| Significance | < $2.2 \mathrm{e}-16$ |  |  | <2.2e-16 |  |  |

Table 8. AFT Model Estimated Results

| AFT Model | One-way |  |  | Two-way |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Number of pedestrians | 1132 observations |  |  | 927 observations |  |  |
| Explanatory Variable | $\beta$ | Std.Err | P-value | $\beta$ | Std.Err | P-value |
| (Intercept) | -0.294 | 0.174 | 0.09 . | 0.528 | 0.192 | 0.006** |
| Sex | - | - | - | -0.0663 | 0.039 | 0.089. |
| Age | 0.199 | 0.0448 | <0.001*** | 0.09 | 0.0413 | 0.029* |
| Distraction | - | - | - | 0.104 | 0.0563 | 0.065 . |
| Group Size | - | - | - | -0.0394 | 0.0148 | 0.008*** |
| Near Side or Far Side | - | - | - | -0.104 | 0.0448 | 0.021* |
| Vehicle Type | 0.111 | 0.0611 | 0.07 . | - | - | - |
| Distance | 0.00131 | 0.0012 | 0.287 | - | - | - |
| (Distance)^2 | $-1.67 \mathrm{E}-05$ | 5.88E-06 | 0.0046* | - | - | - |
| Approaching Speed | 0.024 | 0.00681 | <0.001*** | 0.0142 | 0.00637 | 0.025* |
| (Approaching Speed)^2 | -0.000363 | 0.00022 | 0.1000. | - | - | - |
| Adjacent Vehicle | 0.243 | 0.0602 | <0.001*** | 0.227 | 0.0504 | <0.001*** |
| Close Follower Indicator | 0.214 | 0.059 | $<0.001^{* * *}$ | 0.184 | 0.053 | $<0.001^{* * *}$ |
| Nr Ped. Waiting | 0.052 | 0.0129 | <0.001*** | - | - | - |
| Nr Ped. Crossing | -0.0494 | 0.00765 | <0.001*** | -0.05 | 0.00871 | $<0.001^{* * *}$ |
| Log Likelihood | -1766.8 |  |  | -1581.6 |  |  |
| Degree of freedom | 11 |  |  | 10 |  |  |
| Significance | <2.2e-16 |  |  | <2.2e-16 |  |  |
|  |  |  |  |  |  |  |

The log-linear and AFT models showed similar results. In the AFT framework, the exponential of the estimated coefficient is called the accelerated factor (AF), which measures, for each variable, the increased pedestrian delay associated with an increase in the value of that variable. For example, the exponential of a positive coefficient, such as Age Indicator in the one-way model, is $\mathrm{AF}=\exp (0.199)=$ 1.22, which means that with the increase of Age indicator by 1, it shows an increase of $22 \%$ probability of waiting. Conversely, the exponential of a negative coefficient, such as Group Size, is $\exp (-0.104)=0.901$. The interpretation is that a pedestrian is likely to wait about 0.901 times as long ( $9.9 \%$ shorter) when the group size increases by 1 person, while keeping all the other variables unchanged. Generally, a coefficient greater than zero (or, equivalently, an exponent parameter greater than 1.0) indicates that an increase in the explanatory variable results in increased pedestrian delay, and vice versa.

### 5.2.1. Before and After Studies

### 5.2.1.1. Distance and Speed

In both the regression model and the AFT duration model, the squared term of the vehicle-to-pedestrian distance and the squared term of vehicle speed were highly significant in the one-way case, indicating that there exists a non-monotonic relationship between pedestrian delay and the two variables. Figure 11 shows that pedestrian wait time is greatest when the interacted vehicle is $39-46 \mathrm{ft}$ from the crosswalk, approaching at an average speed. The pedestrian wait time is smaller when the vehicle is closer than $39-46 \mathrm{ft}$, because the pedestrian is content to let the vehicle pass before crossing the street with increases as the speed increases and as the distance decreases only up to certain thresholds, after which the relationship becomes the opposite. Such a non-monotonic relationship (See Figure 11) has not been identified in past studies.

However, only speed term showed significant impact on pedestrian waiting time in two-way operation.


Figure 11.Relationships between Pedestrian Delay and Distance \& Speed (One-Way Case)

### 5.2.1.2. Pedestrian Characteristics

Pedestrian Characteristics (Sex, Distraction) have significant impacts on pedestrian waiting behavior in the two-way case. Male showed significant lower waiting durations than females (See Figure 12). Moreover, distraction (Talking and Cellphone Using) will result in a longer waiting time (See Figure 12).


Figure 12.Effects of Pedestrian Characteristics (Two-Way)

### 5.2.1.3. Environmental Factors

The variables near side and far side showed significant impacts on pedestrian waiting durations in the twoway case. If the interacted vehicle is in far lane, the model indicates that pedestrians will have a lower wait time. Furthermore, with the increase in group size, the subject pedestrian will wait less (See Figure 13).


Figure 13.Effects of Environmental Factors (Two-Way)

## Model Performance

### 5.2.2. Mean Absolute Percentage Error (MAPE)

To compare the model performance between the regression models and the AFT duration models, the mean absolute percentage error (MAPE) is used. The MAPE is a summary measure widely used for evaluating the accuracy of prediction results. It can be calculated using Equation (18).

$$
\begin{equation*}
\text { MAPE }=\frac{1}{n} \sum_{i=1}^{n} \frac{O_{i}-P_{i}}{O_{i}} \tag{18}
\end{equation*}
$$

where $O_{i}$ is the observed waiting duration for the i-th pedestrian, $P_{i}$ is the predicted wait duration for the i -th pedestrian, and n is the number of pedestrians included in the model.

A lower MAPE value indicates a higher accuracy of the prediction model. In this study, the MAPE value was calculated as $47.3 \%$ for the log-linear model and $37.6 \%$ for the log-logistic AFT duration model in the one-way case; $46.6 \%$ for the log-linear model and $36.4 \%$ for the log-logistic AFT duration model in the two-way case. In safety studies related to human factors, the MAPE value range from $21 \%$ to $50 \%$ is at a reasonably accuracy level (Chung, 2010).

### 5.2.3. Error Tolerance

Another measure of model prediction accuracy used in duration modeling is related to a certain tolerance of the actual durations (Chung, 2010; Yang et al., 2015). In this part, we defined the percentage error as the percentage difference between the observed and predicted value. The prediction accuracy under certain error tolerance is calculated as the ratio of the predicted durations with percent errors smaller than the given error tolerance to the total number of prediction points. Figure 14 presents the prediction accuracy under error tolerance from $0 \%$ to $100 \%$ for the two estimation models and in the one-way and two-way cases. The plots show that the log-logistic AFT duration model outperformed the log-linear model at each tolerance level in term of the prediction accuracy for both the one-way and two-way cases.


Figure 14.Prediction Accuracy under Different Error Tolerance

## Summary

In this chapter, we used the first event analysis to estimate the pedestrian waiting behavior. The occurrence of the first event (interaction) is considered to have the most critical impact on pedestrian wait durations. Log-linear model and AFT model were utilized to investigate the effects of covariates on pedestrian wait behavior. Based on the results, parameters distance and speed show non-monotone relationship on pedestrian wait durations in one-way case. The peak values for distance ( 39.4 ft in AFT model and 46.8 ft in Log-Linear) and speed ( $33 \mathrm{ft} / \mathrm{s}$ in AFT model and $37 \mathrm{ft} / \mathrm{s}$ in Log-Linear) are shown to cause the longest delay on pedestrian at semi-controlled crosswalks. For example, when a vehicle too close to yield, it will not necessarily cause the confusion because pedestrian will let vehicle go first then cross. When a vehicle is too far, there's no need for pedestrian to hesitate and no confusion arises. The most confusing part is that, within a certain distance, pedestrian feels unsafe and the car slows down to yield. Both parties are both delayed. These phenomena are common in unsignalized crosswalks and these results have not been investigated by existing literature. In future studies, we should not only consider strategies that can increase the vehicle yielding rate, but also find strategies to reduce the delay in terms of control speed and interacted distance. One typical strategy is to force vehicle brake at a certain distance if the pedestrian flow is really high while if pedestrian flow is not high, then let vehicles cross as normal.

On the two-way street, pedestrian a less likely to wait more when an object vehicle is on the far side suggested by models. Based on our observations, compared with one-way operation, pedestrians have different crossing strategies in two-way street because they have to care vehicles in different lanes from different sides. If the object vehicle is in the far lane, pedestrians are assertive to cross even if the far lane vehicle does not yield. Interactions between pedestrians and far-lane vehicles may result in a lower post encroachment time (PET) so that some dangerous cases are likely to occur in two-way street unsignalized crosswalks. Conflict analysis can be done to examine the results in the future.

## 6. PEDESTRIAN WAIT TIME - MARKOVIAN APPROACH

### 6.1. Model Formulation

Consider a Markov renewal process $\left(J_{n}, T_{n}\right)$, where $\mathrm{T}_{0}<\mathrm{T}_{1}<\ldots<\mathrm{T}_{\mathrm{n}}<\infty$ are the successive times of entry to states $J_{o}, J_{l}, \ldots, J_{n}$. If $S_{n}=T_{n}-T_{n-1}$ is the sojourn time (gap time or lag time), the Markov renewal kernel $Q_{h j}(d)$ is a cumulative distribution function of time:

$$
\begin{align*}
Q_{h j}(d) & =P\left(J_{n+1}=j, S_{n+1} \leq d \mid J_{0}, J_{1}, \ldots, J_{n}=h, S_{1}, S_{2}, \ldots, S_{n}\right)  \tag{19}\\
& =P\left(J_{n+1}=j, S_{n+1} \leq d \mid J_{n}=h\right)
\end{align*}
$$

$J_{0}, J_{l}, \ldots, J_{n}$ is an embedded homogeneous Markov chain taking values in a finite state space with transition probability:

$$
\begin{equation*}
p_{h j}=P\left(J_{n+1}=j \mid J_{n}=h\right)=\lim _{t \rightarrow \infty} Q_{h j}(t), n \in N \tag{20}
\end{equation*}
$$

We define the distribution function of the sojourn time in state $h$ by:

$$
\begin{equation*}
H_{h}=\sum_{j=1}^{s} Q_{h j}(t), \quad \forall t \in \square \tag{21}
\end{equation*}
$$

The probability distribution function of sojourn time (gap time or lag time), through the transition probabilities of the embedded Markov chain in terms of conditional probability, is:

$$
\begin{equation*}
F_{h j}(d)=P\left(S_{n+1} \leq d \mid J_{n+1}=j, J_{n}=h\right)=\frac{Q_{h j}(d)}{p_{h j}} \tag{22}
\end{equation*}
$$

$F_{h j}(d)$ is a cumulative probability distribution and is called a sojourn time in state $h$ if the next state will be $j$. Based on equation (22), we can write the probability density function as $f_{h j}(d)$. The hazard function $\mathrm{H}_{\mathrm{hj}}$ of $F_{h j}(d)$ will be:

$$
\begin{equation*}
\alpha_{h j}=\lim _{\Delta d \rightarrow 0} \frac{\operatorname{Pr}\left(d<S_{n+1} \leq d+\Delta d \mid J_{n+1}=j, J_{n}=h, S_{n+1}>d\right)}{\Delta d}=\frac{f_{h j}(d)}{S_{h j}(d)}=\frac{f_{h j}(d)}{1-F_{h j}(d)} \tag{23}
\end{equation*}
$$

### 6.1.1. Distribution of Durations

For the Semi-Markov process, we need to first assume that sojourn time (gap time or lag time) belongs to a specific parametric distribution. The sojourn time, given any state $h$ to state $j$, is modeled as a random variable from the best fitted distribution. The SemiMarkov package in R software (Listwon-Krol and SaintPierre, 2015) offers three distributions -- Exponential, Weibull and Exponential Weibull. Based on maximum likelihood estimation, the Weibull distribution (Weibull, 1951) was chosen to model the sojourn time from state $h$ to state $j$. The Weibull distribution is defined as the probability density function:

$$
\begin{equation*}
f(x \mid \lambda, k)=\frac{k}{\lambda}\left(\frac{x}{\lambda}\right)^{k-1} e^{-\left(\frac{x}{\lambda}\right)^{k}} \tag{24}
\end{equation*}
$$

According to Equations (23) and (24), the hazard ratio for the Weibull distribution is:

$$
\begin{equation*}
\alpha_{h j}=\frac{f_{h j}(d)}{1-F_{h j}(d)}=\frac{k}{\lambda}\left(\frac{x}{\lambda}\right)^{k-1} \tag{25}
\end{equation*}
$$

Table 9.Weibull Distribution Duration Parameters

| Duration Parameters in Weibull Distribution |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Transition | One-Way |  |  |  | Two-Way |  |  |  |
|  | $\lambda$ |  | k |  | $\lambda$ |  | k |  |
|  | Estim. | SE | Estim. | SE | Estim. | SE | Estim. | SE |
| $1 \rightarrow 2$ | 2.864 | 0.12 | 1.364 | 0.06 | 3.021 | 0.09 | 1.785 | 0.07 |
| $1 \rightarrow 3$ | 2.014 | 0.04 | 1.843 | 0.04 | 2.229 | 0.06 | 1.794 | 0.05 |
| $2 \rightarrow 3$ | 1.881 | 0.11 | 1.329 | 0.07 | 1.67 | 0.11 | 1.075 | 0.05 |

Table 10.Wald Test of Weibull Distribution

| One-Way |  |  | Two-Way |  |  |
| :--- | :--- | :--- | :--- | :--- | :--- |
| Transition | Wald Test | P-value | Transition | Wald Test | P-value |
| $1 \rightarrow 2$ | 43.53 | $<0.0001$ | $1 \rightarrow 2$ | 128.59 | $<0.0001$ |
| $1 \rightarrow 3$ | 357.06 | $<0.0001$ | $1 \rightarrow 3$ | 259.26 | $<0.0001$ |
| $2 \rightarrow 3$ | 19.84 | $<0.0001$ | $2 \rightarrow 3$ | 1.89 | 0.1692 |



Figure 15.Density Functions Between Different Transitions.

Figure 15 shows the probability density function for every transition. The transition 1-2-3 (reject the lag, then accept the next gap) can be expressed as the convolution product $f_{1-2-3}=\int_{0}^{x} f_{1-2}(u) f_{2-3}(x-u) d u$, where $\mathrm{x}=$ the total wait time during the transition 1-2-3. The probability density function $f_{1-2}(u)$ permits the calculation of the probability that the transition from state 1 to state 2 takes place in $\mu$ time units. The probability density function $f_{2-3}(x-u)$ leads to the probability that the transition from state 2 to state 3 takes place in the remaining $x-\mu$ time units. Table 11 shows the most likely wait times for each distribution.

Table 11.The Most Likely Wait Times

| The most likely wait times | One-Way | Two-Way |
| :--- | :--- | :--- |
| Transition 1-2 | 1.087 s | 1.906 s |
| Transition 1-3 | 1.317 s | 1.415 s |
| Transition 1-2-3 | 3.15 s | 3.38 s |

On the one hand, Transition 1-2-3 includes recurrent events -- rejecting the first lag (the first vehicle does not yield), then accepting the following gap (second interacted vehicle yields). Although there is information about the two events, the survival model only used information about the first event (Transition $1-2)$. On the other hand, when a pedestrian experiences Transition 1-2, it means that he/she rejects a gap and his/her accepted wait time is greater. The Semi-Markov model considers state 2 as a right-censored state. The Semi-Markov model has the potential to estimate the pedestrian actual accepted wait time continuously through the Markovian renewal process until the last observation is observed.

### 6.1.2. Parameterization

In order to illustrate the influence of covariates associated with the Semi-Markov process, the Cox proportional model (Cox, 1972) was used. Let $\mathrm{Z}_{h j}$ be a vector of explanatory variables related to the transition from $h$ to $j$ and $\beta_{h, j}$ be the vector of estimated regression parameters. By the Cox proportional model:

$$
\begin{equation*}
\alpha_{h j}\left(d \mid Z_{h, j}\right)=\alpha_{h j}(d) \exp \left(\beta_{h, j}{ }^{T} Z_{h, j}\right) \tag{26}
\end{equation*}
$$

According to Equation (26), we took advantage of transition-specific variables $\mathrm{Z}_{\mathrm{h}, \mathrm{j}}$ defined in Table 6. Using the Semi-Markov package in R, significant factors were retained. The estimated regression coefficients are shown in Table 12. Table 12 illustrates the effects of covariates on the sojourn time in each transition. Positive coefficients denote the increasing risk or accelerating factors. while negative coefficients demonstrate decreasing risk. We will discuss the effects of significant variables in the following sections.

Table 12.Parametric Effects of Multi-State Model

| Coefficients | One-Way |  |  | Two-Way |  |  |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- |
|  | $\mathbf{1 \rightarrow \mathbf { 2 }}$ | $\mathbf{1 \rightarrow \mathbf { 3 }}$ | $\mathbf{2 \rightarrow \mathbf { 3 }}$ | $\mathbf{1 \rightarrow \mathbf { 2 }}$ | $\mathbf{1 \rightarrow \mathbf { 3 }}$ | $\mathbf{2 \rightarrow \mathbf { 3 }}$ |
| Near or Far Side | - | - | - | - | - | - |
| Group Size | - | - | - | 0.123. | $0.078^{* *}$ | $0.199^{*}$ |
| Nr. Ped. Crossing | $-0.319^{* * *}$ | $0.105^{* * *}$ |  | $-0.148^{* * *}$ | $0.082^{* * *}$ | - |
| Close Follower <br> Indicator | $-0.459^{* * *}$ | - | - | -0.184. | $0.196^{*}$ | - |
| Adjacent Vehicle | - | $-0.196^{*}$ | $0.450^{*}$ | - | - | - |
| Sex | - | - | $0.325^{*}$ | - | - | $0.353^{*}$ |
| Age | - | $-0.355^{* * *}$ | $0.312^{*}$ | - | $-0.371^{* * *}$ | - |
| Nr. Ped. Waiting | - | $-0.157^{* * *}$ | - | $0.134^{* *}$ | - | - |
| Hesitation | $-1.364^{*}$ | $-1.287^{* * *}$ | $-1.316^{* * *}$ | -0.351. | $-1.165^{* * *}$ | $-0.877^{* * *}$ |
| Vehicle Type | $-0.173^{* * *}$ | - | - | - | $0.318^{*}$ | - |
| Distance | $-0.635^{* * *}$ | $0.303^{* * *}$ | - | $-0.456^{* * *}$ | - | - |
| Approaching <br> Speed | $0.057^{* * *}$ | $-0.031^{* * *}$ | - | $0.034^{* * *}$ | $-0.011^{*}$ | - |
| Distraction | - | - | - | - | - | -0.312. |
| Log-Likelihood | -1128.8095 vs $-1318.376($ Null LL) | $-1097.4645 \mathrm{vs}-1191.7082$ (Null LL) |  |  |  |  |
| .p<0.1; p<.05; ** p<.01; *** p<.001 |  |  |  |  |  |  |

### 6.1.3. Hazard of Semi-Markov Process

The hazard rate of a Semi-Markov process is defined as the probability of transition towards state $j$ between the time $d$ and $d+\Delta d$, given that the process is in state $h$ for a duration $d$.

$$
\begin{align*}
\gamma_{h j}(d) & =\lim _{\Delta d \rightarrow 0} \frac{\operatorname{Pr}\left(J_{n+1}=j, d<S_{n+1} \leq d+\Delta d \mid J_{n}=h, S_{n+1}>d\right)}{\Delta d} \\
& =\left\{\begin{array}{l}
\frac{q_{h j}}{1-H_{h}(d)}=\frac{p_{h j} f_{h j}(d)}{1-H_{h}(d)}=\frac{p_{h j}\left(1-F_{h j}(d)\right) \alpha_{h j}}{1-H_{h}(d)} \text { if } p_{h j}>0 \text { and } H_{h}(d)<1 \\
0 \quad \text { otherwise }
\end{array}\right. \tag{27}
\end{align*}
$$

Note that Equation (25) and Equation (27) demonstrate two different hazards. To better understand the differences, we defined the hazard in Equation (25) as the hazard given transition from state $h$ to state $j$. The hazard defined by Equation (27) is the hazard of the Semi-Markov process, which represents the immediate probability of going to state $j$ given state $h$ in a small-time interval [ $\mathrm{d}, \mathrm{d}+\Delta \mathrm{d}$ ] (Dominicis and Manca, 1984). Therefore, for the state space $I=\{1,2,3\}$, we can use Equation (27) to calculate the "staying" probability for the case $h=j$ :

$$
\begin{align*}
& p_{11}(d)=e^{-\int_{0}^{d} \gamma_{12}(\tau)+\gamma_{13}(\tau) d \tau} \\
& p_{22}(d)=e^{-\int_{0}^{d} \gamma_{23}(\tau) d \tau}  \tag{28}\\
& p_{33}(d)=1
\end{align*}
$$

Consequently, we can calculate the probabilities of each transition in a Markov chain as:

$$
\begin{align*}
& p_{1-2}(d)=\int_{0}^{d} p_{11}(\tau) \gamma_{1-2}(\tau) p_{22}(d-\tau) d \tau \\
& p_{1-3}(d)=\int_{0}^{d} p_{11}(\tau) \gamma_{1-3}(\tau) p_{33}(d-\tau) d \tau \\
& p_{2-3}(d)=\int_{0}^{d} p_{22}(\tau) \gamma_{2-3}(\tau) p_{33}(d-\tau) d \tau  \tag{29}\\
& p_{1-2-3}(d)=\int_{0}^{d} p_{11}(\tau) \gamma_{1-2}(\tau) p_{2-3}(d-\tau) d \tau
\end{align*}
$$

An interpretation will be needed to illustrate the transition $P_{1-2-3}(d)$ in Equation (29). The term $p_{11}(\tau) \gamma_{1-2}(\tau)$ denotes the transition from state 1 to 2 (pedestrian rejects the first lag, or the first vehicle doesn't yield) in $\tau$ duration time. $p_{2-3}(d-\tau)$ indicates the probability of transferring from state 2 to state 3 (pedestrian accepts the next gap) in the remaining time $d-\tau$.
$P_{1-3}$ and $P_{1-2-3}$ in Equation (29) are the total waiting behavior of pedestrians in the curb area, because the total number of transitions 1-3+1-2-3 is 966 out of 1132 for the one-way case. Therefore, we can use the transition probability of the Semi-Markov process $P_{l-3}+P_{l-2-3}$ to explain the variables in Table 12 .

### 6.2. Model Discussions

For a one-way operation:

### 6.2.1. Number of Pedestrians Impacts

(A). Figure 16 (a) shows that the number of pedestrians already on the crosswalk will speed up the Transition 1-3 ( $-0.319^{* * *}$ in Table 12). Pedestrians on a crosswalk is an indication that it is safe to cross. Nevertheless, the number of pedestrians on crosswalks will cause much more delay for Transition 1-2-3 ( $0.105^{* * *}$ in Table 12 for Transition 1-2). This indicates that, if there are many pedestrians already on the crosswalk and a pedestrian chooses to wait, he or she should wait for a longer time for 1-2-3.
(B). Figure 16 (b) indicates that the number of pedestrians waiting on curb will result in a delay for subject pedestrian to cross for Transition 1-3 (-0.157*** in Table 12).


Figure 16. Pedestrian Impacts

## Vehicle Dynamics

Figure 17(a) demonstrates that a platoon of vehicles has little impact on the decision-making process for Transition 1-3 because pedestrians only "negotiate" with the leading vehicle while making a decision. The close follower indicator has effects on delay on Transition 1-2-3 (-0.459*** for Transition 1-2 in Table 12), however. For Transition 1-2-3, if a pedestrian chooses to wait for a platoon of vehicles, he or she can expect to wait for a longer time compared with individual vehicle.

The adjacent vehicle indicator has a negative impact on Transition 1-3 (-0.196* in Table 12) because pedestrians have to "negotiate" with two different vehicles in different lanes. See Figure 17(b).


Figure 17. Multiple Vehicle Effects

### 6.2.2. Hesitation

The Hesitation parameter (-1.364* for Transition 1-2;-1.286*** for Transition 1-3; and -1.316*** for Transition 2-3 in Table 12) has effects on delay on pedestrian waiting behavior. This is intuitive because, if a pedestrian hesitates in the curb area, the misunderstanding between pedestrians and motorists increases. This will delay the pedestrian's time to cross. See Figure 18.


Figure 18. Hesitation Parameter Effects

### 6.2.3. Pedestrian Characteristics

Figure 19 illustrates the effects of pedestrian characteristics on pedestrian waiting process. Compared with other groups, young pedestrians are more assertive for Transition 1-3 ( $-0.355 * * *$ in Table 12) because they have lower wait durations than do other groups. See Figure 19(a). Besides, males are less likely than females to wait for subsequent gaps ( $0.325^{*}$ in Table 12). See Figure 19(b).

(a) Age

Transition 1-3


## Figure 19. Pedestrian Characteristics Impact

### 6.2.4. Distance and Speed

(A). Distance has significant impact on pedestrian waiting behavior. For Transition 1-3 (0.303*** in Table 12), with the increase of distance to vehicle, pedestrian wait time is reduced. However, it has inverse effects on the transition 1-2-3. A shorter distance to the first interacted vehicle will result in a faster transition 1-2 $\left(-0.635^{* * *}\right.$ in Table 12) and then speed up the waiting process for Transition 1-2-3. The closer the vehicle is to the crosswalk, the more uncertain and unsafe a pedestrian feels. "Let the car go first" will be a safe crossing strategy for pedestrians, if the interacted vehicle is too close to yield. See Figure 20.


Figure 20. Distance Parameter Effects
(B). The effects of vehicle speed on pedestrian decision making are shown in the following figure. We can see that, with the increase in speed, the pedestrian wait time of Transition 1-3 ( $-0.031 * * *$ in Table 12) is increasing, while the pedestrian wait time of transition 1-2-3 is decreasing. The first interacted vehicle will pass the area quickly with a higher speed, which result in a faster transition 1-2 ( $0.057 * * *$ in Table 12) so that it reduces the delay for transition 1-2-3. "Let the car go first" will also be a safe crossing strategy for pedestrian to cross, if the interacted vehicle is too fast to yield. See Figure 21.


Figure 21. Speed Parameter Effects

### 6.3. Before and After Studies

### 6.3.1. Group Effects

The Markovian model shows results similar to the Survival model. The number of pedestrians in a group significantly decreases the pedestrian delay in transitions 1-3 (0.078** in Table 12) and 2-3 (0.199* in Table 12). See Figure 22.


Figure 22. Group Effects on Two-Way Case

### 6.3.2. Using Cellphone or Talking

While using a cellphone or talking, pedestrians have a lower risk moving from state 2 to state 3 ( -0.312 . in Table 12), which results in a higher delay. See Figure 23.


Figure 23. Distraction Effects on Two-Way Case

### 6.3.3. Vehicle Type

Vehicle Type affects delay in Transition 1-2-3 with one-way street operation period. This indicates that, if a pedestrian chooses to yield to a bus (Transition 1-2 -0.173*** in Table 12), he/she has to wait longer. However, after conversion to a two-way street, the crossing probability for Transition 1-3. This is reflected in a higher probability to choose Transition 1-3 (0.318* in Table 12) with two-way operation. After the conversion from one-way to two-way operation, Lafayette CityBus was required to remove most bus routes from University Street. Therefore, in the two-way case, there were fewer buses.

### 6.3.4. Adjacent Vehicle

Adjacent vehicle has no effects on two-way cases. This means that whether a pedestrian waits or not is less likely to be affected by an adjacent vehicle when a pedestrian-motorist interaction occurs. We found the similar results in the analysis of motorist behavior.

## 7. CONCLUSIONS AND RECOMMENDATIONS

### 7.1. Pedestrian-Motorist Interaction

Table 13 summarizes the variables (defined in Table 1) that were found to be significant in explaining the pedestrian and motorist behavior seen in the recorded video at the University Street crosswalks. Some variables were found to be significant in both the one-way and two-way street cases. Combining the oneway and two-way data, we were able to identify factors that changed significantly between the two cases. These are shown with check marks in the "Significant difference" columns of Table 13.

Table 13.Summary of Model Results

| Variable | Pedestrian Mo |  | Motorist Mod |  |
| :---: | :---: | :---: | :---: | :---: |
|  | Significant in both cases | Significant difference | Significant in both cases | Significant difference |
| ApprSpeed | $\checkmark$ | - | $\checkmark$ | - |
| Hesitation | $\checkmark$ | $\checkmark$ | $\checkmark$ | $\checkmark$ |
| FlowWith/On | - | $\checkmark$ | $\checkmark$ | - |
| SlowsDown | $\checkmark$ | - | - | - |
| Distance | $\checkmark$ | - | - | $\checkmark$ |
| AdjVeh | - | $\checkmark$ | - | $\checkmark$ |
| NoF | $\checkmark$ | - | - | $\checkmark$ |
| GroupSize | - | $\checkmark$ | - | - |
|  |  |  |  |  |
| - = not applicable |  |  |  |  |

The model results are consistent with expectations in terms of the direction of influence. Some examples are described below.

A pedestrian is more likely to cross during an interaction at the semi-controlled crosswalk if the approaching vehicle is moving at a slow speed, is slowing down, or is far enough away from the crosswalk. However, what this research offers is a more quantitative assessment of the pedestrian response to these and other factors, as well as highlighting the importance of a pedestrian's actions with respect to hesitation. While the findings of the pedestrian model are largely behavioral, there are some practical aspects to the
findings. For example, the approach speed finding above can be translated into a speed that would lead to a desired likelihood of pedestrians choosing to cross, all else being equal.

The study found factors that affect an approaching driver's behavior, which focused on a driver's likelihood of slowing down for pedestrians, rather than the likelihood of yielding. Examples of these findings with respect to two-way vehicle traffic operation are:
A. A greater number of pedestrian characteristics factors (Hesitation, Distraction and FlowOn) have a significant impact on a driver's willingness to decelerate in the two-way case than in one-way operation.
B. Except for the speed variable (AppSpeed), variables concerning vehicle dynamics and characteristics become insignificant (CloseFollow and AdjVeh), when compared with one-way operation. This means that a driver on the two-way street is less likely to be affected by a closefollowing vehicle or by an adjacent vehicle when a pedestrian-motorist interaction occurs.
C. Environmental characteristics factors (Distance and NoF) became significant in a driver's decision to slow down, compared to one-way operation.

A driver's decision is mainly influenced by interacted pedestrian behavior and the environmental characteristics when an interaction occurs. The change of one-way to two-way operation removed the effects of interaction between vehicles (CloseFollow and AdjVeh) on a driver's decision and led drivers to react more to the interacted pedestrian.

### 7.2. Pedestrian Waiting Time

This research describes the data and models used to analyze the wait durations of pedestrians when they interact with vehicles at "semi-controlled" crosswalks. The variables for 2059 pedestrian wait durations were carefully defined and measured from video recordings.

Survival models and multi-state Markov models were developed and compared. In Survival models, time-to-first-event analysis was conducted, which suggested a non-monotonic relationship of distance and speed on pedestrian waiting behavior in one-way operation. In multi-state Markov models, recurrent events analysis was undertaken, which examined transition-specific covariates on pedestrian waiting behavior. Both models showed different results.

Survival models consider the first event as the most important interaction and suggested different waiting behaviors in terms of pedestrian characteristics, vehicle dynamics and environmental factors when one-way converted to two-way operation. Multi-state semi-Markov model suggested consistent pedestrian waiting behaviors when one-way convert into a two-way operation. Compared with survival models,

1. Multi-state Markov models can better explain the covariates impact in different waiting transitions (1-3 vs 1-2-3) and the impact of vehicle yielding on pedestrian wait durations, which provides indepth insights about pedestrian wait strategies. For example, after rejecting the first lag (the first vehicle doesn't yield and second vehicle comes), males are more unwilling to wait than females.
2. Semi-Markov models help to explain the non-monotone relationship of speed and distance found in survival model.
3. Multi-state Markov models can be further improved by including random effects to correlate ingroup observations. Furthermore, it can be flexible if more sojourn time distributions are tested.
4. One limitation for Markovian model is that existing statistical software cannot deal with the Transition 2-2. This lost the information when pedestrians rejected multiple gaps. We hope to solve this problem in my future research.

In Chapter 5, Survival models reveal a non-monotone relationship between distance, speed and pedestrian wait time. In Chapter 6, we analyze the effects of vehicle yielding behavior on pedestrian wait durations which provides information for non-monotone relationships. The non-monotone relationship is a paradox that if a vehicle too close and too fast to yield, the strategy - "let vehicle go first" will not cause much delay for pedestrian. However, if vehicle is too close and too fast, many dangerous interactions will probably occur. If the vehicle is neither too far nor too close, the game between pedestrian and motorist will be more complicated. There is a tradeoff for city engineer to consider. On the one hand, use control strategies such as bumper or pedestrian signals to increase vehicle yielding rates while this will result in a higher vehicle delay and pedestrian confusion in specific areas. On the other hand, encouraging drivers to "ignore" far side pedestrians when the interacted distance is close so that it will increase the efficiency. However, this will may cause a lower post encroachment time (PET) and higher crash rates. Besides, future research should focus specifically on the areas where most complicated games between pedestrians and motorists occur. Simple model structures may not explain the complicated games and some of the most advanced frontier of modeling approaches such as Markov switching models, multivariate models etc. can be further explored. Finally, both survival models and Markovian models reveal that pedestrian in a group will result in a lower delay for pedestrian in two-way case. City engineers can consider the geometry designs near the crosswalks to indirectly increase pedestrian arriving groups when pedestrian flow is high.

## 8. SYNOPSIS OF PERFORMANCE INDICATORS

### 8.1 Part I

One (1) transportation-related course was offered during the study period that was taught by the PI and a teaching assistant who are associated with the research project. One (1) graduate student participated in the research project during the study period. One (1) transportation-related advanced degree (doctoral) program utilized the CCAT grant funds from this research project, during the study period to support graduate students. One student supported by this grant received a doctoral degree. This research project was leveraged to obtain $\$ 109,709$ in additional funding for projects such as SPR-4301: Assessment of an Offset Pedestrian Crossing for Multilane Arterials.

### 8.2 Part II

Research Performance Indicators: 3 journal articles and 3 conference articles were produced from this project. The research from this advanced research project was disseminated to over 200 people from industry, government, and academia. The research was presented at several conferences, including the 2019 Purdue Road School in West Lafayette, the 2019 Next Generation Transportation Systems Conference (NGTS), the 2019 ITE (Purdue Chapter) Annual Dinner, the 2018 TRB annual meeting, and the 2019 TRB annual meeting. The outputs, outcomes, and impacts are described in the following sections.

## 9. OUTPUTS, OUTCOMES, AND IMPACTS

### 9.1. Research Outputs

### 9.1.1. Synopsis of Project

The project exceeded expectations, in that the results went far beyond a basic "inventory" and categorization of interactions between pedestrians and motorists. Appropriate statistical analysis has revealed factors and relationships that are described in three papers listed in the following section.

### 9.1.2. List of Publications

Fricker, J. D., \& Zhang, Y. (2019). Modeling Pedestrian and Motorist Interaction at Semi-Controlled Crosswalks: The Effects of a Change from One-Way to Two-Way Street Operation. Transportation Research Record. https://doi.org/10.1177/0361198119850142.

Zhang, Y., Qiao, Y., \& Fricker, J. D. (2020). Investigating Pedestrian Waiting Time at Semi-Controlled Crossing Locations: Application of Multi-State Models for Recurrent Events Analysis. Accident Analysis \& Prevention, 137, 105437.
https://www.sciencedirect.com/science/article/pii/S0001457519308759

Zhang, Y., \& Fricker, J. D. (2020). Multi-State Semi-Markov Modeling of Recurrent Events: Estimating Driver Waiting Time at Semi-Controlled Crosswalks. Analytic Methods in Accident Research, 100131. https://www.sciencedirect.com/science/article/pii/S221366572030021X

### 9.1.3. List of Presentations

Jon, D. Fricker, Yunchang Zhang (2019) Modeling Pedestrian and Motorist Behavior at Semi-Controlled Crosswalks: The Effect of a Change from One-Way to Two-Way Street Operation. Presented at the Transportation Research Board 98th Annual Meeting, January 2019.

Yunchang Zhang, Jon, D. Fricker (2019). "Smart Interaction - Pedestrians and vehicles in a CAV environment". Presented at the 1st Annual Conference on Next-Generation Transport Systems, May 2019.

Yunchang Zhang, Jon, D. Fricker (2020). "Multi-State Semi-Markov Models: An Application to Drivers' Gap Acceptance in front of Approaching Pedestrians at Unsignalized Crosswalks". Presented at the Transportation Research Board 99th Annual Meeting, January 2020.

### 9.1.4. List of Outcomes and Highlights

The outcomes and highlights from this project are:

- "Semi-controlled" crosswalks are unsignalized but marked with "yield to pedestrian" signs.
- Pedestrians and motorists engage in non-verbal "negotiation" to decide priority.
- Video recordings were made of 2059 pedestrians interacting with 1003 motorists.
- A conversion from 1-way to 2-way operation allowed a before-and-after study at the same location.
- The probabilities of pedestrian wait time are quantified under alternative scenarios.


### 9.1.5. List of Impacts

This study improves the operation and safety of semi-controlled crosswalks by developing a database and identifying factors that affect pedestrian and motorist behavior.

1. This information can be used to test the impact of new technologies on crosswalk safety and performance.
2. We maintain an ongoing relationship with the Area Plan Commission of Tippecanoe County, whose staff is engaged in data collection of pedestrian, bicycle, and scooter activity at busy locations downtown and near campus. We shared data and analysis of "hotspots".
3. A coupling project with INDOT is a perfect complement to this study, in that it offers opportunities to apply a variety of designs and control methods to other types of crossing locations.

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# APPENDIX: JOURNAL PAPERS PUBLISHED FROM THIS WORK 

CCAT Project Title: Pedestrian-Vehicle Interaction in a CAV Environment: Explanatory Metrics

Fricker, J.D., and Zhang, Y. (2019). Modeling Pedestrian and Motorist Interaction at Semi-Controlled Crosswalks: The Effects of a Change from One-Way to Two-Way Street Operation, Transportation Research Record 2673(11), 433-446.


#### Abstract

: A large number of crosswalks are indicated by pavement markings and signs, but are not signalcontrolled. In this paper, such a location is called "semi-controlled." At locations where such a crosswalk has moderate amounts of pedestrian and vehicle traffic, pedestrians and motorists often engage in a nonverbal "negotiation" to determine who should proceed first. This paper describes the detailed analysis of video recordings of more than 3,400 pedestrian-motorist interactions at semi-controlled crosswalks. The study also took advantage of a conversion from one-way operation in spring 2017 to two-way operation in spring 2018 on the street chosen for data collection and analysis. This permitted before and after studies at the same location. The pedestrian models used mixed effects logistic regression and binary logistic regression to identify factors that influence the likelihood of a pedestrian crossing under specified conditions. The complementary motorist models used generalized ordered logistic regression to identify factors that impact a driver's likelihood of decelerating, which was found to be a more useful factor than likelihood of yielding to pedestrian. The data showed that $56.5 \%$ of drivers slowed down or stopped for pedestrians on the one-way street. This value rose to $63.9 \%$ on the same street after it had been converted to two-way operation. Moreover, two-way operation eliminated the effects of the presence of other vehicles on driver behavior. Relationships were found that can lead to policies and control strategies designed to improve the operation of such a crosswalk. Video recordings were made of pedestrians using crosswalks at locations in West Lafayette, Indiana, where "State Law Yield to Pedestrian Within Crosswalk" signs (Figure 1a) were present. Pedestrians at such crossings theoretically have priority over approaching vehicles in Indiana. Observations indicate that confusion exists among pedestrians and motorists, however, because the sign's message is subject to varying interpretations. Sometimes drivers stop and let pedestrians standing at the curb cross the street, and sometimes drivers fail to yield to pedestrians entering the crosswalk. Although marked crosswalks can have relatively high fatality rates per pedestrian crossing, no safety issues were observed during the data collection at the locations studied; it was primarily a matter of delay.


Zhang, Y. and Fricker, J.D (2020). Multi-state semi-Markov modeling of recurrent events: Estimating driver waiting time at semi-controlled crosswalks, Analytic Methods in Accident Research, 28(1), 100131.


#### Abstract

: At "semi-controlled" crosswalks, signs and markings are present, but delay to pedestrians and motorists is largely the result of the "negotiation" between the two parties to determine who yields. This paper proposes a novel approach using multi-state semi-Markov models to investigate motorists' delay and their interactions with pedestrians. Motorist waiting behavior can be divided into a series of gap acceptance decisions as part of a Markov Chain. Each gap acceptance decision is modeled as a specific transition between two states in the Markov Chain. To demonstrate the reliability of the proposed models, multi-state semi-Markov models are estimated for the waiting behavior of more than 1,000 drivers in the presence of pedestrians at semi-controlled crosswalks. The multi-state semi-Markov models are capable of dealing with specific challenges related to (i) the need to account for recurrent events and (ii) a generalized framework for vehicle delay estimation and simulation at semi-controlled crosswalks. The extent to which motorists behave more aggressively and impatiently as their delay increases is demonstrated. Differences in behavior for operators of buses and trucks were also identified. The semiMarkov method is also able to deal effectively with the "pulsing" arrival patterns of pedestrians at crosswalks as university classes begin and end nearby and handle temporal heterogeneity. Finally, to address aggressive driver behavior, several safety implications are discussed.


Zhang, Y., and Fricker, J.D. (2021). Incorporating conflict risks in pedestrian-motorist interactions: A game theoretical approach, Accident Analysis \& Prevention, 159(1), 106254.


#### Abstract

: At "semi-controlled" crosswalks with yield signs and markings, negotiations as to the right-of-way occur frequently between pedestrians and motorists, to determine who should proceed first. This kind of "negotiation" often leads to traffic delay and potential conflicts. To minimize misunderstandings between pedestrian and motorist that can have serious safety consequences, it is essential that we understand the decision-making process as the "players" interact in real street-crossing situations. This paper employs a game-theoretic approach to investigate the joint behaviors of pedestrians and motorists from the perspective of safety. Assuming bounded rationality for each player, the quantal response equilibrium is a special kind of game with incomplete information. Explanatory variables such as conflicting risks and time savings can be incorporated into the payoff functions of the "players" via expected utility functions. Finally, model parameters can be estimated using an expectation maximization algorithm. The gametheoretic framework is applied to model pedestrian-motorist interactions at a semi-controlled crosswalk on a university campus. The estimation results indicate that the likelihood of pedestrian-vehicle conflict can be quantified. The results can lead to control measures that facilitate the negotiation between pedestrian and motorist and reduce the conflict risk at semi-controlled crosswalks.


