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DOT HS-802 425

A METHODOLOGY FOR ANALYZING GENERAL CATEGORICAL DATA WITH MISCLASSIFICATION ERRORS WITH AN APPLICATION IN STUDYING SEAT BELT EFFECTIVNESS

Contract No. DOT-HS-4-00897 June 1977 Final Report

PREPARED FOR:

U.S. DEPARTMENT OF TRANSPORTATION NATIONAL HIGHWAY TRAFFIC SAFETY ADMINISTRATION WASHINGTON, D.C. 20590

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Technical Report Documentation Page

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| 1. Report No. | 2. Government Access | ion No. | 3. Recipient's Catalog N | ło. | | | |
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| 12. Sponsoring Agency Name and Address | | | Final Danamt | | | | |
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ACKNOWLEDGMENTS

Sincere appreciation is expressed to those individuals and organizations who supplied the data for this study. In addition to the many accident victims who responded to the survey, the author wishes to thank the North Carolina Division of Motor Vehicles for providing the sample of accident records, and the following 11 hospitals that assisted in the data gathering:

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Alamance County Hospital, Burlington, N.C. Cabarrus Memorial Hospital, Concord, N.C. Cape Fear Memorial Hospital, Wilmington, N.C. Forsyth Memorial Hospital, Winston Salem, N.C. Gaston Memorial Hospital, Gastonia, N.C. High Point Memorial Hospital, High Point, N.C. King's Mountain Hospital, Kings Mountain, N.C. Memorial Hospital of Alamance County, Burlington, N.C. Onslow Memorial Hospital, Jacksonville, N.C. Wake County Memorial Hospital, Raleigh, N.C. Watts Hospital, Durham, N.C.

The author is also very indebted to Ms. Lucy Smith, Research Assistant at HSRC, who was responsible for the coordination of the collection of the supplementary data. As such, she directed the entire effort involved in the hospital and telephone surveys. Special thanks are also due to Ms. Debbie Wood, a programmer at HSRC, whose contributions both in writing original programs and in using a large number of existing ones are evident throughout this work.

A large number of HSRC staff participated in various parts of this project, especially in connection with the supplementary sample. These include Dr. B.J. Campbell, HSRC Director, who contributed throughout the study with advice as to basic issues regarding the hospital and telephone surveys; Ms. Jane Youngblood, who played a major role in conducting the telephone survey; Ms. Anita Leung, who carried out some additional

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required programming; and Ms. Jane Stutts, who greatly assisted in the preparation of this final report.

A special acknowledgment is due to the participation of Dr. Donald W. Reinfurt, Project Director, in all parts of this work. His advice, guidance, and organizational assistance, as well as his critical reading and suggestions for improving the content and layout of this report, were essential to the final product.

Finally, the author is greatly indebted to Mr. Donald F. Mela, the Contract Technical Manager, for his helpful guidance and suggestions through all phases of the project.

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I. INTRODUCTION

There are several major problems which make a precise measure of seat belt effectiveness very difficult. One of these is the presence of misclassification errors in police-reported accident data. To date, most studies on the effectiveness of seat belts in reducing injury have been based on police-level data. Due to the circumstances surrounding the officer's investigation of the crash, however, this data generally contains misclassification errors relating to belt usage and injury sustained. Such errors have the potential of seriously biasing any effectiveness estimates derived from that data.

The problem of misclassification errors in police-reported accident information with reference to studies of seat belt effectiveness was first raised by Mela (1974) and further discussed in Hochberg (1976). The discussion in the latter report supported the need for a methodology for modeling and obtaining unbiased inferences from general categorical data with misclassification errors.

Much has been written on the effects of misclassification errors on studies of association in 2 × 2 contingency tables (see, e.g., Fleiss, 1973, Ch. 3). In Koch (1969), the misclassification errors are assumed to be generated according to a random response error model. As such, the methodology is based on repeated classifications of the experimental elements. Such a methodology, however, can not always be satisfactory because of obvious practical difficulties and since, in many problems, misclassification errors are fixed biases rather than random response errors. Most studies of the potential effects of fixed bias misclassification errors have severely restricted the number of error parameters examined. In the 2 \times 2 table setup, one may theoretically have as many as 12 different parameters for fixed bias misclassification errors. (For example, an element which actually belongs in the first row and the second column may be misclassified into the second row and second column, etc.). In practice, however, many of these parameters are assumed to be zero. Thus, Bross (1954) introduced a model for fixed bias misclassification errors for a 2 \times 2 table where only two error parameters are considered; and Hochberg (1976) discussed the effects of six error parameters on three measures of association in 2 \times 2 tables of belt usage by level of injury.

While the effects of more general misclassification error structures on inference have been discussed in some recent works (see, e.g., Goldberg, 1975), no methodologies for an improved statistical inference have been presented. The purpose of the present report is to present such a methodology, and to apply it to the study of safety belt effectiveness.

The setup for the methodology is general; i.e., the discussion can be applied to any multidimensional cross-classified data obtained by unrestricted random sampling. The methodology itself is based primarily on the double sampling scheme originally introduced by Tenenbein (1970, 1971, 1972) for estimating the parameters of a multinomial classification when misclassification errors prevail.

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The following situation is assumed. There are two classification "devices" available. One is expensive to apply, yet gives "correct" results, while the other is relatively inexpensive but "fallible." As

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an example, Diamond and Lilienfeld (1962) discuss an experimental situation in public health research where the true classification device is a physician's examination, whereas the fallible classifier is a questionnaire completed by the patient. In other situations, the "true device" and "fallible device" may simply refer to making or not making an extra effort to obtain more reliable data.

The methodology, as developed in this report and applied to the study of belt effectiveness, uses an original large sample based on (fallible) police-reported data, and requires that a small subsample of the data be cross-classified by means of some "true" classifying device. In this case, the true classifier is assumed to be hospital reports on the injured occupants and telephone interviews for the non-injured occupants. The supplementary sample of cross-classified data is then used to adjust the original police-based sample, and inference of seat belt effectiveness is taken from this larger, <u>adjusted</u> sample (the adjusted police data).

In real world problems, it is often the case that the true classification device uses different nominal scales than those used by the fallible device. This is illustrated by the seat belt data presented in this report, where injury is coded by the police using the K,A,B,C,O scale, but reported by the hospitals according to the Abbreviated Injury Scale (AIS). In such instances, use of the two-sample methodology has the additional advantage of enabling one to carry out an efficient study expressing results in terms of the (often) finer scale utilized with the relatively small supplementary sample. Thus, for the present study, final estimates of seat belt effectiveness could be based on the AIS scale, rather than the less precise K,A,B,C,O scale.

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In summary, the two-sample methodology proposed in this report represents one approach to resolving the problems of inference arising from classification errors in categorical data. To highway safety researchers concerned with the issue of safety belt effectiveness, it is offered as a viable alternative to drawing inference solely from police report data (which may be biased) or obtaining an independent reliable sample of sufficient size and basing inference entirely on it (a process that is likely to be both costly and time consuming).

The methodology itself is described briefly in Chapter II and in detail in Appendix A. In Chapter III, the original large data source and the supplementary data used to demonstrate the technique are described. Chapter IV presents the results of applying the technique to this data, while Chapter V presents a general discussion of the nature and actual magnitude of misclassification errors in the data. This chapter is the outgrowth of an effort to test certain hypotheses concerning misclassification errors that were made in the Hochberg (1976) report. Finally, Chapter VI provides a discussion of the methodology with suggestions for further research in this area.

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II. METHODOLOGY

The statistical methodology developed in this research (see Appendix A for details) pertains to the setup where all variables are subjected to misclassification errors when the fallible device (i.e., police reports) is used. It is assumed herein that the magnitudes of errors within combinations of levels of the correctly reported set of variables are possibly different.

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The procedure basically extends Tenenbein's (1970, 1971, 1972) double sampling scheme originally introduced for estimating the parameters of a multinomial distribution when misclassification errors prevail. The procedure simultaneously utilizes the information from a large "fallible" sample (in this case, a large collection of police-reported data) along with a relatively small "non-fallible" supplementary sample (in this case, data from telephone and emergency room respondents) to more efficiently estimate the multinomial parameters ($\pi(\underline{i})$) of interest (namely, belt usage by injury category).

The cross-classification of the resulting data by both police reports and non-police reports results in contingency tables with underlying multinomial distributions. The task is to find efficient estimators for the parameters of the resulting distributions along with covariances of these estimators for subsequent hypothesis testing.

The details of the estimation and testing procedures along with the necessary notation are given in Appendix A. In a nutshell, the procedure consists of two stages. In the first stage, Maximum Likelihood techniques are utilized to estimate the overall true distribution of occupants in accidents across the levels of This is done by setting up the joint likelihood function (A.1) for the <u>combined</u> sample, differentiating, and setting the partial derivatives equal to zero. This yields the MLE's given in (A.2) which are related to the main parameters of interest $(\pi(i))$ by (A.3).

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Asymptotic covariances of the $\hat{\pi}$'s are next derived by Taylor series expansions. For efficiency, these estimates serve as initial input to the asymptotically equivalent Least Squares procedures presented in Grizzle et al. (1969) for additional inferences concerning linear hypotheses involving the $\hat{\pi}$'s.

Thus, the estimates make use of the information in both samples to derive estimates of the π 's. If the supplementary (or non-fallible) sample were sufficiently large, it would be optimal to use only this sample. However, the procedure suggested herein allows for a relatively small but expensive "non-fallible" sample supplemented by the large but inexpensive "fallible" sample to carry out improved estimation and test-ing procedures for such categorical data problems.

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III. THE DATA

The data used to demonstrate and evaluate the methodology presented in Chapter II was derived from North Carolina traffic accidents. The original large sample consisted of over 139,000 occupants in accidents involving cars or small trucks for which police report information on belt usage was available. The accidents were those recorded on the HSRC North Carolina accident tapes for the first eight months of 1975. Table 3.1 presents the raw data, broken down by age and sex of driver, car "make" and model year, vehicle damage severity, and accident type. (The "not stated" or "unknown" categories are deleted from the table.)

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In order to not only examine the effect of misclassification errors in this police-reported data but also to adjust the belt effectiveness estimates accordingly, supplementary data was obtained for a subsample of over 2,000 North Carolina accidents. For this phase of the study, it was assumed that follow-up telephone interviews would provide "true" information on belt usage and injury level for the non-injured occupants, while special forms completed by participating hospitals would supply the corresponding "correct" information on the injured occupants. During the four month data collection period, over 2100 telephone interviews were successfully completed, and over 900 hospital forms linked with accident reports.

Appendix B presents the cross-classification of the police and nonpolice data by belt usage and injury level across a number of variables indicated on the accident report form. These include age and sex of driver, car "make" and model year, vehicle damage severity (TAD), and accident type.

Table 3.1 Belt usage and injury level for 1975 North Carolina accident data, broken down by age, sex, car "make", model year, vehicle damage severity (TAD) and accident type

| | | | No Injury | | | Injury | ······································ | |
|-------------------|---------------|----------------|-----------------|------------------------|----------------|-----------------|--|--------|
| | | No Belt (U) | Lap Belt (L) | Lap & Sh. Belt (LS) | No Belt (U) | Lap Belt (L) | Lap & Sh. Belt (LS) | Total* |
| Over | rall | 94834 | 11287 | 3006 | 21127 | 2010 | 493 | 132757 |
| Але | 16-55 | 80777 | 9579 | 2703 | 18316 | 1712 | 440 | 113527 |
| | 56+ | 11460 | 1404 | 225 | 2225 | 237 | 40 | 15591 |
| Sex | Male | 63095 | 7817 | 2077 | 12417 | 1194 | 286 | 86886 |
| | Female | 29729 | 3232 | 868 | 8253 | 773 | 197 | 43052 |
| Car "Make" | U.S. | 52139 | 6502 | 1664 | 11546 | 1074 | 241 | 73166 |
| | Foreign | 5369 | 691 | 428 | 1700 | 189 | 74 | 8451 |
| | 1960-1968 | 26594 | 1634 | 68 | 6296 | 264 | 7 | 34863 |
| Model | 1969-1971 | 19553 | 2159 | 243 | 4268 | 384 | 38 | 26645 |
| Year | 1972-1973 | 15950 | 3052 | 363 | 3239 | 521 | 69 | 23194 |
| | 1974-1975 | 8363 | 1472 | 1527 | 1748 | 276 | 214 | 13600 |
| Vehicle Damage | Minor | 36931 | 4355 | 1140 | 3876 | 372 | 85 | 46759 |
| Severity | Severe | 21958 | 2530 | 745 | 10443 | 918 | 224 | 36818 |
| Accident | Non-Collision | 11758 | 1159 | 388 | 7100 | 457 | 129 | 20991 |
| Туре | Collision | 81150 | 9897 | 2560 | 13593 | 1510 | 352 | 109062 |

* "Not Stated" cases excluded.

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The remaining sections of this chapter describe in greater detail the processes involved in obtaining the supplementary telephone and hospital data, along with the weighting of the supplementary sample required to make it representative of the overall accident sample.

The Telephone Survey

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For this phase of the study, the North Carolina Division of Motor Vehicles (DMV) furnished HSRC copies of randomly selected accident report forms recently received from across the state. The initial quota of 250 accident reports per week was gradually increased to 450 per week to build up the sample size for the uninjured occupants. In addition, some 300 supplemental accident reports were obtained from the local police departments in Chapel Hill and Raleigh.

As copies of the accident reports (see Appendix.^C) were received by HSRC, they were screened to exclude injured occupants as well as accidents involving motorcycles, pedestrians, tractor trailers, etc. Next, as the North Carolina accident report form does <u>not</u> provide the telephone numbers of the drivers involved in the accident, a rather complicated and time-consuming, yet educational process was carried out to locate these individuals. This involved first trying to obtain an appropriate telephone number and then reaching the desired person for the interview. Some of the difficulties encountered included the following:

(a) The names and telephone numbers of females were particularly difficult to locate in the telephone directories since wives are usually not listed separately from their husbands, nor daughters from their fathers. In the case of married women, this problem could have been alleviated somewhat had the police also recorded the husband's name (rather than "same as driver") under the vehicle ownership heading on the accident report. Consequently, considerable effort was devoted to

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examining the telephone directories to try and match last names with street addresses, or, for the smaller towns, imposing on the telephone operator to perform this task.

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- (b) In North Carolina, there are a large number of small, private telephone companies operating throughout the state. Not too infrequently, two different telephone companies operated within a given radius of a community, but an individual's telephone number was naturally only listed in the directory of the company that owned his phone. Thus, one could not conclude that a given telephone number was not available after looking in just one directory, unless one was certain that there was not a second telephone company operating in the area. Unfortunately, this information was not always available and thus the "no listing" frequency was inflated.
- (c) North Carolina has two different area codes (919 in the East, 704 in the West) with no available listing of which towns are in which area. This only further served to complicate the job of the information operator for addressees in some of the smaller municipalities.

As the telephone numbers became available, the interviewers concentrated their calling during the early evening hours, primarily on weekdays. Initial contact with family members did, however, often require follow-up calls of the accident-involved occupant during office hours, late in the evening, or on weekends. All telephone numbers were attempted for at least three days before being classified as "not reachable".

Appendix D contains a copy of the questionnaire that was utilized, along with the suggested introductory remarks to be used by the interviewer. Over the four-month interviewing period, for nearly half of the accident-involved drivers there was no telephone listing and hence these people were unfortunately not reachable. For an additional few cases, the subject was not available for interview. Of the remaining cases, only 2.9 percent of those contacted flatly refused to cooperate, while an additional 0.5 percent denied being in an accident. The upshot was questionnaire information on 2,132 uninjured occupants, along with the corresponding accident information from the accident report forms.

The Hospital Survey

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The second data source used to evaluate the accuracy of policereported belt usage and injury level required the cooperation of hospitals and, in particular, their emergency room staffs. More specifically, it required that the hospitals submit a completed form (see Appendix E) with information on belt usage and degree of injury for each patient seen in the emergency room as a result of an automobile accident.

Fourteen hospitals across North Carolina were contacted as potential participants in this phase of the study. All but one of the hospitals had assisted during 1972 in a similar type study (see McLean, 1973). The following eleven hospitals participated in this study resulting in statewide information on "correct" belt usage and injury level:

> Alamance County Hospital, Burlington, N.C. Cabarrus Memorial Hospital, Concord, N.C. Cape Fear Memorial Hospital, Wilmington, N.C. Forsyth Memorial Hospital, Winston Salem, N.C. Gaston Memorial Hospital, Gastonia, N.C. High Point Memorial Hospital, High Point, N.C. King's Mountain Hospital, Kings Mountain, N.C. Memorial Hospital of Alamance County, Burlington, N.C. Onslow Memorial Hospital, Jacksonville, N.C. Wake County Memorial Hospital, Raleigh, N.C. Watts Hospital, Durham, N.C.

After each hospital administration agreed to assist in the data collection, a training session with its emergency room staff was held. The keystone of the training session was the stressing of the importance of inquiring about seat belt usage while the patient was being treated. With those unconscious or disoriented cases, the emergency room staffs were encouraged to question witnesses or even the ambulance services personnel regarding belt usage. Because of their expressed interest in

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the outcome of the survey, all hospitals were promised (and will receive) a copy of the completed report.

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Appendix E contains a copy of the form completed by the emergency room staffs on each accident victim. The form (HSR-006) was designed with the goal of being comprehensive and yet easy to complete in the midst of emergency-type pressures and confusion. For ease and accuracy in making the correct injury (AIS) classification, the standard American Medical Association scale definitions on injury categories were incorporated into the form (see page 67). The hospital staff were further instructed to call HSRC (collect) if they had any questions regarding either the forms or some broader aspect of the study. Bi-monthly newsletters were also issued by HSRC to clear up any problem areas as well as to offer encouragement and support to the participating personnel. As a result, there were no major difficulties associated with this phase of the study.

The hospitals collected data on accident victims from March 1, 1975 through June 1, 1975. As forms were completed, they were mailed to HSRC in the pre-addressed business reply envelopes provided. Each week, HSRC compiled a list of the <u>name</u> of each injured occupant reported by the cooperating hospitals along with his <u>birthdate</u>, <u>county</u> of residence, and <u>date(s)</u> of treatment, and then forwarded this information to the Division of Motor Vehicles (DMV). The DMV staff then, to the extent possible, located the accident reports corresponding to the names on the list.

Due to time delays in receiving accident reports from the various police agencies across the state, there were inevitable difficulties in

locating the corresponding accident report forms. Two weeks into the study, DMV requested that in addition the driver's name be submitted with each accident victim's name. It was anticipated that this would increase the percentage of linkages with the accident records file. Through the regular newsletter to the hospitals, this additional step was quickly implemented by the hospital personnel.

Even with this additional information with which to link the emergency room data with the police accident data, it was not always possible to locate the corresponding copy of the police report form. This was as anticipated due to occasional lengthy delays in DMV receiving the reports from some of the smaller or more remote police departments. Also, if the accident victim provides false information (names) in the emergency room setting, that case will not likely be able to be used. Nevertheless, the rate of linkage (slightly over 70%) appeared reasonable and no serious biases were evident.

Once a hospital report form was linked with its corresponding accident report from the DMV file, information from both sources was keypunched and placed on file for subsequent analysis. As mentioned previously, a total of 911 emergency room forms were successfully linked and coded during this phase of the project.

Adjustment of the Supplementary Sample

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Clearly, the supplementary sample described in the preceding two sections is not a simple random sample from the larger sample of North Carolina accidents for which police-reported information is available. Actually, it is structured as a stratified random sample where its two strata are based generally on injury level. With very few

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exceptions, those occupants interviewed over the phone did indeed place lower on the injury scale than those for whom information was obtained via the hospital reports.

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In order to account for any biases that such a sampling scheme might introduce, the data were adjusted to reflect the overall target population (i.e., all North Carolina accident victims) with respect to certain relevant variables, namely those whose distributions are confounded in the design--age, sex, race, and level of injury.

Table 3.2 compares the police-reported data for the supplementary sample <u>only</u> with <u>all</u> of the police-reported data for the first half of 1975, prior to any adjustment (i.e., weighting).

| Variable | Level | Supplementary Sample | 1975 Accidents |
|-----------------|---|------------------------------------|----------------------------------|
| Age | 16-55 | 87.0 | 87.8 |
| | 56+ | 13.0 | 12.2 |
| Sex | Male | 60.7 | 66.8 |
| | Female | 39.3 | 33.2 |
| Race | White | 83.9 | 77.1 |
| | Non-White | 16.1 | 22.9 |
| Injury Level | No Injury C Injury B Injury A Injury Fatality | 74.0 10.6 11.9 3.2 0.3 | 82.7 7.7 7.0 2.3 0.3 |

Table 3.2 Age, sex, race, and injury distributions of supplementary sample and 1975 accident data.

As expected, the unadjusted supplementary sample inflates the proportions of injury, except for fatalities. It also oversamples whites and females. Again, this might be anticipated, since these individuals

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are, for example, generally easier to contact in a telephone survey. Finally, the age deviations between the two samples are small.

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As a result of this investigation, the supplementary sample was weighted to match the relevant 1975 accident data with respect to its distribution over the 40 cells in the cross-classification of (age) \times (sex) \times (race) \times (injury level).

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IV. RESULTS

In this chapter are presented estimates of injury risk and belt effectiveness based on the North Carolina police-reported data only, the supplementary (non-police) data only, and the combined police and supplementary data (applying the methodology described in detail in Appendix A). This is done for a number of control variables of interest, including age and sex of driver, model year and type of car, vehicle damage severity (TAD), and accident type.

The specific procedure utilized is the modified Maximum Likelihood approach described in Appendix A. This approach is the most convenient to apply, and the results are generally equivalent to those obtained via the complete Least Squares approach.

Due to the relatively small size of the supplementary sample, only two levels of each control variable were considered. These are defined as follows:

| Sex: | Male Female |
|------|----------------|
| Age: | 16-55 56+ |

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Car type: U.S. (e.g., Chevrolet, Plymouth) Foreign (e.g., VW Beetle, Datsun)

Model year: pre-1972 1972-75

Vehicle damage severity:

Minor (i.e., front center or front left
 impacts of TAD severity 1; all other
 impacts with severity ≤ 3)

Moderate or severe (i.e., front center or front left impacts of TAD severity > 1; all other impacts with severity > 3)

Accident type:

"Collision" (e.g., collision of motor vehicle in road with another motor vehicle, pedestrian, bicyclist, etc.)

"Non-Collision" (e.g., ran-off-road on the right, overturn)

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As previously noted, in the police-reported data, "injured" is defined as having an injury of "C" or worse on the K,A,B,C,O scale, while for the supplementary data, a person recorded as injured has an AIS severity rating of 1 or greater.

Table 4.1 gives the risk of injury and the corresponding belt effectiveness estimates derived from the <u>police-reported</u> data, Table 4.2 for the <u>supplementary</u> (or non-police) data, and Table 4.3 for the <u>combined</u> sample. The measures of belt effectiveness (E) are presented for two cases -- E_{12} for none vs. lap, and E_{23} for lap vs. lap and shoulder -where

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Thus, belt effectiveness is viewed as the percentage decrease in injury as one becomes progressively more restrained.

| | | Percent | Injured | (K,A,B,C) | Belt Effectiveness | |
|----------|-------------------------|------------------------------|-----------------|-----------------|--------------------|--------------------|
| Lont | rol Variable | None | Lap | Lap & Sh | None vs Lap | Lap vs Lap & Sh |
| Sex | Male | 16.44 (0.13) ¹ | 13.25 (0.36) | 12.10 (0.67) | 19.42 (2.27) | 8.66 (5.63) |
| | Female | 21.73 (0.21) | 19.30 (0.62) | 18.50 (1.19) | 11.17 (3.00) | 4.16 (6.90) |
| Age | 16-55 | 18.48 (0.12) | 15.16 (0.34) | 14,00 (0.62) | 17.97 (1.91) | 7.67 (4.57) |
| | 56+ | 16.26 (0.32) | 14.44 (0.87) | 15.09 (2.20) | 11.17 (5.61) | -4.51 (16.47) |
| Model | 1960-71 | 18.63 (0.16) | 14.59 (0.53) | 12.64 (1.76) | 21.67 (2.93) | 13.37 (12.47) |
| 1 cui | 1972-75 | 17.02 (0.22) | 14.98 (0.49) | 13.02 (0.72) | 12.00 (3.09) | 13.05 (5.59) |
| Accident | Collision | 14.35 (0.11) | 13.24 (0.32) | 12.09 (0.60) | 7.74 (2.33) | 8.68 (5.06) |
| 1990 | Non-Collision | 37.65 (0.35) | 28.28 (1.12) | 24.95 (1.90) | 24.89 (3.06) | 11.77 (7.58) |
| Car | U.S. | 18.13 (0.15) | 14.18 (0.40) | 12.65 (0.76) | 21.81 (2.31) | 10.76 (5.94) |
| туре | Foreign | 24.05 (0.51) | 21.48 (1.38) | 14.74 (1.58) | 10.69 (6.06) | 31.36 (8.59) |
| Damage | Minor | 9.50 (0.15) | 7.87 (0.39) | 6.94 (0.73) | 17.15 (4.31) | 11.83 (10.22) |
| Jevenicy | Severe | 32.23 (0.26) | 26.62 (0.75) | 23.12 (1.35) | 17.40 (2.43) | 13.17 (5.65) |
| Tot | tal Sample ² | 115961 | 13297 | 3499 | | |

Table 4.1 Estimated injury risks and belt effectiveness, based on <u>police-reported</u> data.

¹Standard deviation

²"Not Stated" cases excluded

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| , | | Percent | Injured | (AIS ≥ 1) | Belt Eff | ectiveness |
|---------------------------|---------------|------------------------------|-----------------|-----------------|------------------|--------------------|
| Cont | trol Variable | None | Lap | Lap & Sh | None vs Lap | Lap vs Lap & Sh |
| Sex | Male | 26.31 (1.30) ¹ | 20.42 (2.39) | 14.46 (2.73) | 22.38 (9.87) | 29.21 (15.73) |
| | Female | 38.09 (1.92) | 27.72 (4.45) | 23.64 (5.73) | 27.21 (12.26) | 14.74 (24.79) |
| Age | 16-55 | 31.68 (1.17) | 22.67 (2.26) | 17.28 (2.74) | 28.43 (7.60) | 23.80 (14.25) |
| | 56+ | 21.33 (2.82) | 19.51 (6.19) | 13.79 (6.40) | 8.51 (31.44) | 29.31 (39.75) |
| Model Year | 1960-71 | 31.13 (1.37) | 23.88 (3.01) | 11.36 (4.78) | 23.29 (10.23) | 52.41 (20.91) |
| | 1972-75 | 28.53 (1.81) | 20.99 (3.03) | 18.29 (2.92) | 26.40 (11.59) | 12.90 (18.75) |
| Accident | Collision | 27.71 (1.19) | 21.73 (2.33) | 14.12 (2.67) | 21.60 (9.06) | 35.02 (14.13) |
| туре | Non-Collision | 41.71 (2.59) | 25.76 (5.38) | 22.73 (6.32) | 38.25 (13.46) | 11.77 (30.69) |
| Car | U.S. | 29.82 (1.14) | 21.61 (2.21) | 16.88 (2.96) | 27.52 (7.91) | 21.93 (15.85) |
| гуре | Foreign | 36.84 (3.69) | 27.78 (7.47) | 16.95 (4.88) | 24.60 (21.62) | 38.98 (24.04) |
| Damage | Minor | 18.86 (1.41) | 15.17 (2.69) | 12.50 (3.38) | 19.59 (15.46) | 17.59 (26.62) |
| Severity | Severe | 42.36 (1.82) | 36.36 (4.37) | 25.00 (4.51) | 14.15 (10.96) | 31.25 (14.92) |
| Total Sample ² | | 1783 | 384 | 218 | | |

Table 4.2 Estimated injury risks and belt effectiveness, based on <u>supplementary</u> data.

¹Standard deviation ²"Not Stated" cases excluded ين

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| , | | | ercent Inju | ured | Belt Effectiveness | | |
|----------|-------------------------|------------------------------|--------------------------|-----------------|--------------------|--------------------|--|
| Conti | rol Variable | None | Lap | Lap & Sh | None vs. Lap | Lap vs Lap & Sh | |
| Sex | Male | 26.99 (1.02) ¹ | 20.17 (2.26) | 14.02 (2.71) | 25.26 (9.02) | 30.48 (15.96) | |
| | Female | 38.58 (1.55) | 28.88 (4.12) | 23.10 (5.64) | 25.16 (11.22) | 20.01 (22.63) | |
| Age | 16-55 | 32.03 (0.93) | 22.81 (2 . 13) | 16.60 (2.70) | 28.79 (7.06) | 27.23 (14.01) | |
| | 56+ | 23.00 (2.27) | 19.97 (5.38) | 15.09 (6.73) | 13.16 (26.19) | 24.46 (40.01) | |
| Mode1 | 1960-71 | 31.17 (1.07) | 22.77 (2.78) | 12.04 (4.88) | 26.96 (9.40) | 47.11 (22.60) | |
| 'ear | 1972-75 | 29.98 (1.47) | 21.68 (2.83) | 17.91 (2.94) | 27.69 (10.30) | 17.40 (18.03) | |
| Accident | Collision | 27.79 (0.97) | 22.27 (2.19) | 13.11 (2.52) | 19.84 (8.53) | 41.13 (12.94) | |
| Туре | Non-Collision | 46.22 (1.89) | 25.97 (4.87) | 29.40 (7.88) | 43.81 (10.82) | -13.20 (38.35) | |
| Car | U.S. | 30.50 (0.90) | 21.33 (2.05) | 16.26 (2.86) | 30.09 (7.11) | 23.77 (15.66) | |
| Туре | Foreign | 38.99 (3.20) | 29.77 (7.48) | 16.79 (5.05) | 23.64 (20.69) | 43.62 (22.37) | |
| Damage | Minor | 19.65 (1.21) | 15.24 (2.67) | 11.42 (3.24) | 22.41 (14.72) | 25.06 (25.47) | |
| Severity | Severe | 45.38 (1.41) | 38.08 (4.06) | 27.50 (4.78) | 16.08 (9.45) | 27.79 (15.15) | |
| Tot | tal Sample ² | 115961 | 13297 | 3499 | | | |

Table 4.3 Estimated injury risks and belt effectiveness based on combined police and supplementary data.

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¹Standard deviation ²"Not Stated" cases excluded

In comparing the results based on the police data (Table 4.1) with those based on the combined data (Table 4.3), one finds that the estimated risks and effectiveness are quite different, with the combined estimates being substantially higher in most cases. Thus, for example, controlling for sex, the estimates of belt effectiveness for males are 25.26% for none vs lap and 30.48% for lap vs lap + shoulder, based on the combined results, compared with only 19.42% and 8.66% for the corresponding estimates based on the police data. A part of this difference can be attributed to the lack of equivalence between the two injury scales employed, with fewer people being classified as injured on the police scale. (Perhaps "injured" on the AIS scale should have been defined as AIS $\geq 2!$) Most of the difference, however, is indeed probably due to misclassification errors in the police reports.

Since all the data in this study are biased, their quality (i.e., accuracy) cannot be judged solely on the basis of the accompanying standard deviations. A more appropriate measure of the accuracy of the police report estimates is the more general "mean square error" (MSE), where

 $MSE = Variance + (Bias)^2$.

This measure can be applied to both the risk and effectiveness estimates.

In calculating MSE's for the police report data, we assume that the best available estimator for the bias of a given estimate is the difference between that estimate and the corresponding estimate obtained via the "combined" methodology. For example, the police estimate of percentage injury to unbelted males is 16.44 with a variance of 0.017 and the "combined" sample estimate is 26.99 with a variance of 1.04. The bias

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of the police report estimate is then 10.55 (= 26.99 - 16.44). Thus, the MSE is given by

$$(10.55)^2 + .017 = 111.32$$

This compares with a variance (and approximate MSE) of 1.04 for the "combined" approach.

It should be noted that, in some cases, the MSE's of the police estimates are lower than those of the "combined" approach, even though the estimated biases in the police estimates are quite substantial. This is primarily due to the "relatively" small size of the supplementary sample. If the estimated bias remained the same, but a larger (say, threefold) sample size were available, then all of the estimates based on the combined approach would have much smaller MSE's than those based on only the police-reported data.

While the injury risk and effectiveness estimates based on the police data only differ substantially from the "combined methodology" estimates, a comparison of the supplementary vs. combined data results reveals a generally high level of agreement. This is not unexpected, since both represent consistent estimators, based on the same definition of injury (AIS \geq 1). The positive effect of wearing lap belts and the additional benefit derived from the use of shoulder belts are clearly evidenced in all but a few isolated instances.

In further comparing the supplementary and combined results, one finds that the standard deviations (STD's) of the estimates for the combined approach are usually lower than those resulting from the supplementary data only. However, they are not as much lower as might be anticipated, considering the large increase in sample size. This seems to

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indicate that the combined methodology may offer only a slight improvement in accuracy over using only the supplementary data, at least when multi-dimensional contingency tables are considered with this considerable amount of data.

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By increasing the size of the supplementary sample, one would, of course, decrease the level of error in both the supplementary and combined data results. However, it is suggested that the relative decrease would probably be less for the combined than the supplementary results. That is, as the supplementary sample size is increased, the relative benefit of utilizing the combined samples to estimate the fallible margin in the cross-classified sample will decrease. Such a conclusion should not affect the overall usefulness of the combined methodology, however, since this approach is designed for use in situations where an original large (but fallible) sample is readily available, but where only a relatively small supplementary (non-fallible) sample can reasonably be obtained.

Another important issue is whether the estimates obtained using the combined methodology are the "best" estimates (i.e., those with lowest STD's) obtainable, using only the data available. The answer is that one can probably derive better estimates, even without increasing sample size. If only the supplementary sample is considered, one could use well-known techniques for building models that smooth the original proportions by removing non-significant variations in a multi-factor set-up. The result would be estimates with lower STD's. How to accomplish the same objective with the "combined approach" is a more complicated matter, and will be discussed further in Chapter VI.

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V. THE NATURE OF MISCLASSIFICATION ERRORS IN POLICE-REPORTED BELT USAGE AND INJURY LEVEL

The investigation described herein is aimed at exploring the magnitude of the misclassification errors of belt usage and level of injury in actual statewide police data. Table 9 in Hochberg (1976) shows the effect on several measures of belt effectiveness (namely, RIDIT, relative risk, and the odds ratio) for various magnitudes of a combination of misclassification errors. Certain questions remain: How large are these misclassification errors in actual data? Are the simplifying assumptions made in Hochberg (1976) valid assumptions? Do the magnitudes of the errors depend upon other factors such as age and sex of the driver?

This chapter explores questions such as these using the data reported in the text where the police (P) data are "fallible" while the "true" belt/injury status is given by the hospital/telephone interview (\bar{P}) data. As before, belt (B) includes the use of any restraint system (lap, lap and shoulder) with \bar{B} indicating no restraint used; injury (I) includes any injury (K,A,B, or C) recorded by the police or an AIS ≥ 1 for the non-police data, while \bar{I} indicates no injury.

Table 5.1 gives the raw frequency data for belt usage and level of injury, cross-classified by the police (P) and non-police (\overline{P}) sources.

| Police | | Injured | | Not | Injured | Total | |
|------------|---------|--------------|----------------|----------------|-----------------|--------------|--|
| Non-Police | | Belt No Belt | | Belt | No Belt | | |
| Injured | Belt | 36 | 16 | 33 | 37 | 122 (5.1%) | |
| | No Belt | 6 | 305 | 5 | 227 | 543 (22.7%) | |
| Not | Belt | 6 | 6 | 256 | 216 | 484 (20.3%) | |
| Injured | No Belt | 2 | 27 | 15 | 1194 | 1238 (51.9%) | |
| | Total | 50 (2.1%) | 354 (14.8%) | 309 (12.9%) | 1674 (70.1%) | 2387 | |

Table 5.1 Cross-classification of supplementary sample according to belt status and level of injury.

It is evident from Table 5.1 that the police are much less likely to report that a seat belt was worn and also much less likely to report that an injury was involved than the non-police source.

More specifically, assuming that the non-police (\bar{P}) reports are the "true" classification mechanisms, then, for <u>injured</u> occupants, the police underreported belt use by 43.4% ($=\frac{16+37}{122} \times 100$) and overreported use by only 2.0% ($=\frac{6+5}{543} \times 100$). For <u>uninjured</u> occupants, the respective estimates are 45.9% underreported and 1.4% overreported.

Conversely, for <u>belted</u> occupants, the police underestimated injury by 57.4% ($=\frac{33+37}{122} \times 100$) and overestimated injury by 2.5% ($=\frac{6+6}{484} \times 100$). For <u>unbelted</u> occupants, the corresponding estimates are 42.7% underestimated and 2.3% overestimated.

Clearly this tendency to underreport belt usage <u>and</u> injury level will affect any derived estimates of belt effectiveness. If one defines belt effectiveness as the percentage decrease in risk of injury resulting from wearing a safety belt, i.e.,

then, based on the police-reported data, safety belts have a 20.2% effectiveness. This compares with an effectiveness estimate of 34.0% based on the "true" hospital/telephone data. Thus, due to misclassification errors in the data, the police estimates apparently substantially underrate the effectiveness of safety belts in reducing the likelihood of injury. (The extent to which the hospital/telephone data represent the true situation is, of course, unknown, but is believed to be much closer to reality!)

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As indicated in Hochberg (1976), there are a total of 12 independent misclassification errors that can arise when classifying individuals into the 2 × 2 table of belt usage (B, \overline{B}) by injury level (I, \overline{I}). For example, an individual that is actually belted and injured (B,I) could be incorrectly classified as (\overline{B} ,I), (B, \overline{I}) or (\overline{B} , \overline{I}). From Table 5.1, the "true" number of belted and injured drivers is 122 whereas the reports classified 16 individuals as being unbelted and injured, 33 as being belted and uninjured and 37 as being unbelted and uninjured!

In order to examine further the nature of such misclassification errors in police belt usage and injury data, estimates of the 12 misclassification error probabilities were obtained, along with corresponding estimates of their covariance matrix, following the approach described in Grizzle, Starmer, and Koch (1969). With these estimates, a variety of hypotheses were then tested. Of particular interest were the two "simplifying" assumptions regarding police misclassification errors that were made in Hochberg (1976). These were:

- (i) The probability of a double misclassification error (i.e., both on belt use and injury level) is well approximated by the product of the two marginal error probabilities.
- (ii) The probability of misclassifying an uninjured occupant (either belted or unbelted) as injured is "unlikely".

The results of the corresponding tests of hypotheses are summarized in Table 5.2 along with estimates \hat{p} of the corresponding misclassification errors based on the data presented in Table 5.1, with the non-police classification representing the "true" condition. It is evident from this data that neither of the two basic assumptions ((i) or (ii)) is tenable.

As a final dimension to this analysis of the magnitude and effect of misclassification errors in this police-reported data, an overall

| p | d.f. | x ² |
|--------|---|--|
| . 3265 | 4 | 52.46 |
| .0238 | 4 | 42.00 |
| .0248 | 2 | 12.31 |
| .0234 | 2 | 29.70 |
| .4541 | 4 | 572.46 |
| .5738 | 2 | 167.94 |
| .4273 | 2 | 404.51 |
| .0157 | 4 | 28.46 |
| .0203 | 2 | 11.23 |
| .0137 | 2 | 17.24 |
| .4538 | 4 | 500.54 |
| .4344 | 2 | 93.13 |
| .4587 | . 2 | 407.42 |
| | p . 3265 . 0238 . 0248 . 0234 . 4541 . 5738 . 4273 . 0157 . 0203 . 0137 . 4538 . 4344 . 4587 | p d.f. .3265 4 .0238 4 .0238 4 .0238 2 .0248 2 .0234 2 .4541 4 .5738 2 .4573 2 .0157 4 .0203 2 .0137 2 .4538 4 .4344 2 .4587 2 |

Table 5.2. Hypothesis tests regarding the various misclassification errors in a 2 × 2 table of belt usage vs injury level.

Note: Corresponding p-values are all \leq .005.

Maximum Likelihood Model was fit to the data to examine the dependence of these errors on driver sex (see Table 5.3). The results of the investigation of other factors replacing sex are not detailed herein because it was unfortunately found that the misclassification errors <u>did</u> depend on the levels of these other factors considered (e.g., model year, car size accident type, driver age). Thus, the models did not simplify in a clear-cut manner.

| Source of Variation | d.f.` | x ² | p-value |
|--|-------|----------------|---------|
| Independence of misclassification errors on sex | 12 | 3.16 | > 0.99 |
| Interactions: | | | |
| PI × PB | 1 | 1.80 | > 0.10 |
| (PB,PI) × PI | 1 | 1.02 | > 0.30 |
| (PB,PI) × P̄B | ٦ | 1.05 | 0.30 |
| (P̃B,P̃I) × PI | 1 | 1.71 | > 0.10 |
| (P̃B,P̃I) × PB | 1 | 1.70 | > 0.10 |
| (P̄B,P̄I) × Sex | 1 | 1.20 | > 0.20 |
| $(\overline{P}B \times \overline{P}I \times PB \times PI)$ | 1 | 0.77 | > 0.30 |
| Total error | 19 | 12.41 | >0.80 |
| Equiprobable model | 31 | 765.84 | |

Table 5.3 Analysis of variance table for the Maximum Likelihood Model including sex.

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VI. DISCUSSION

In Chapter IV, injury risks and safety belt effectiveness estimates were presented for the North Carolina police data only, for the supplementary data only, and for the combined two samples. The variables of interest were age, sex, model year and type of car, damage severity, and accident type. The combined sample results were shown to be the most accurate, although their STD's were not a great deal lower than those associated with the supplementary data only.

The original plan for analyzing the data included considering several levels of some of the key variables (such as TAD and model year), and thus to examine the effects of these variables simultaneously. It was also anticipated that injury level could be defined so as to distinguish between the serious and non-serious and the fatal and non-fatal injuries (in addition to injury-no injury).

However, it soon became obvious that, due to the relatively small size of the supplementary sample and the state of the methodology described in Appendix A, the data could only be meaningfully analyzed using a single variable breakdown and the injury-no injury classification, as presented in Tables 4.1 - 4.3. This is due to the requirement that, in order to use the combined methodology outlined in this report, there must be at least one observation for each level of (injury × police) by (level of belt usage × police) by (level of factor under consideration). This requirement was not able to be satisfied except in the single variable framework for this data set.

The results of this analysis suggest that lap belts alone substantially reduce the likelihood of injury and that lap and shoulder belts
together <u>further reduce</u> this likelihood. However, it should be noted that the specific estimates presented are far from satisfactory, due to their large STD's. Also, due primarily to the large STD's, significant differences in belt effectiveness could not be detected between the two levels of any of the factors considered.

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The supplementary sample size used in this report to illustrate and test the two sample methodology was somewhat over 2000 cases. As suggested in Chapter III, it took considerable effort and coordination to collect the additional telephone and hospital data for even this "small" a sample. Nevertheless, it now appears that, in order to make statistically significant statements on safety belt effectiveness using this technique, one should probably have had a supplementary sample three or four times as large.

Increasing the sample size to, say, 10,000 would have positive effects beyond decreasing the STD's. For example, it would enable one to simultaneously study the effects of several factors associated with seat belt effectiveness. It would also permit one to examine the effectiveness of safety belts for other occupants besides the driver.

Two additional caveats should be made regarding the supplementary sample used in the present study. First, the sources of supplementary information regarding seat belt usage and/or level of injury were follow-up telephone interviews for the non-injured drivers and hospital reports for the injured. Whether or not these sources did indeed provide "true" information was not examined.

Second, as noted in Chapter III, the combined sample of drivers interviewed over the phone and those reported by participating hospitals

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was not totally representative of the overall population of North Carolina drivers involved in accidents. While the supplementary sample was adjusted to resemble the overall sample, the effect of these adjustments on the variances and covariances of the estimates was not taken into account in the analysis.

As a result of these limitations and the relatively small size of the supplementary sample referred to earlier, the data in the present study should best be regarded as a mechanism for demonstrating a new technique, rather than as a definitive estimate of safety belt effectiveness. In order to obtain more accurate and reliable results in future applications of this methodology to the study of seat belt effectiveness, one must ascertain that:

- 1. The supplementary sample is sufficiently large.
- 2. The quality of the "true" classifier is examined and proven reliable.
- 3. The supplementary sample is shown to be representative, or if not representative and adjustments are made, these adjustments are accounted for in the statistical analysis.

Finally, it should be noted that, while increasing the size of the supplementary sample will improve the accuracy of the belt effectiveness estimates based on the two-sample methodology, additional research is needed to further improve upon the technique. More specifically, research is needed to incorporate smoothing models for the entries in the supplementary sample, based on relatively few parameters for the misclassification errors. The methodology as it now stands does not allow for using model-predicted estimates of the frequencies in the supplementary cross-classified sample prior to "merging it statistically" with the original sample.

In addition, it is very reasonable to expect that the very large number of misclassification errors (that introduce too many degrees of freedom in the procedures described) could be structured by an appropriate statistical model, resulting in lower STD's for the predicted frequencies. These investigations might well be worth pursuing in the future.

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APPENDIX A

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A Methodology for Analyzing General Categorical Data with Misclassification Errors

The methodology outlined in this appendix pertains to the setup where all variables are subjected to misclassification errors when the fallible device is used. In practice one might come across situations in which only a subset of the variables is subjected to errors. Two cases are of interest. One case is when the magnitudes of errors within combinations of levels of the correctly reported set of variables are possibly different, and the other case is when these errors can be assumed to be the same across the corresponding levels. The examples in this report demonstrate the former case, while considerable research is still needed in order to treat the latter case.

The Setup and Notation

Two independent samples are drawn from the target population. Each is an unrestricted simple random sample. If the actual frame for the population is finite, we adhere to the concept of a 'super' population (see, e.g., Hartley and Sielken, 1975). The first sample of n_1 elements is classified only by the fallible device. Let $i_1' = (i_1', \ldots, i_d')$ index a specific combination of levels of the d variables under study. The second sample of n_2 elements is simultaneously classified by both the false and the true devices. Here again, we use i_1' to index the fallible cell. To index the true classification we use $i_2 = (i_1, \ldots, i_d)$. Also, let $i_m = 1, \ldots, I_m$ and $i_m' = 1, \ldots, I_m'$, $m = 1, \ldots, d$, with $I_1 \times I_2 \times$ $\ldots \times I_d = k$ and $I_1' \times I_2' \times \ldots \times I_d' = k'$.

Next, we introduce notation for the frequencies and parameters in the two samples. To simplify matters, we use the same letters to

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indicate similar conceptual quantities in both samples. The distinction, however, is easily made since the second sample will always have two indices corresponding to true and false classifications, respectively. Thus, $n(\underline{j}')$ denotes the frequency in the \underline{j}' -th cell as obtained in the first sample by the false classifier. Similarly, $n(\underline{i},\underline{i}')$ denotes the frequency in the second sample classified in the \underline{j} -th cell by the true classifier and in the \underline{i}' -th cell on the false categorical scale. Likewise, let $\gamma(\underline{j}')$ and $\gamma(\underline{i},\underline{j}')$ denote the corresponding population proportions. We introduce $\gamma(\underline{j}|\underline{j}') = \gamma(\underline{i},\underline{j}')/\gamma(\underline{j}')$, which is the fraction of times an element actually belongs to cell \underline{j} when reported to be in cell \underline{j}' by the fallible classifier. In addition, the convention of replacing an index by a period to indicate that summation has been taken over that index will be used throughout, e.g., $n(\underline{j}, \cdot) = \sum_{\underline{j}} n(\underline{i}, \underline{j}')$.

The intermediate parameters of interest are clearly the $\gamma(\underline{j}, \cdot)$ for which we use the special notation $\gamma(\underline{j}, \cdot) \equiv \pi(\underline{j})$.

Throughout this work we will use the convention of putting a tilde to indicate a vector. An indexed vector will be used only for the $\gamma(\underline{j}|\underline{j}')$ where $\gamma(\underline{j}') = \{\gamma(\underline{j}|\underline{j}'), \text{ for all } \underline{j}\}$.

Inference Based on Maximum Likelihood Estimates (MLE) of the $\pi(i)$

Given the data, the likelihood function of the $\gamma(\underline{i}|\underline{i}')$ and the $\gamma(\underline{i}|\underline{i}')$ is given by the following:

$$F = g \prod_{i} [\gamma(i')]^{n(i')+n(\cdot,i')} \prod_{i} b(i') \prod_{i} [\gamma(i|i')]^{n(i,i')}, \quad (A.1)$$

where g is a constant depending on the n_i , i = 1,2, the $n(\underline{i}')$, and the $n(\cdot,\underline{i}')$; the $b(\underline{i}')$ are constants depending on the $n(\cdot,\underline{i}')$ and the

$$\widehat{\gamma}(\underline{i}|\underline{i}') = \frac{n(\underline{i},\underline{i}')}{n(\cdot,\underline{i}')}$$

$$\widehat{\gamma}(\underline{i}') = \frac{n(\underline{i}') + n(\cdot,\underline{i}')}{n_1 + n_2} .$$
(A.2)

Since the $\pi(\underline{j})$ and the $\gamma(\underline{j}|\underline{j}')$ are in 1:1 relation with the set of $\gamma(\underline{j}')$ and $\gamma(\underline{j}|\underline{j}')$, the MLE's of the $\pi(\underline{j})$ are given by

$$\widehat{\pi}(\underline{i}) = \sum_{\underline{i}'} \widehat{\gamma}(\underline{i}') \widehat{\gamma}(\underline{i} | \underline{i}'), \forall \underline{i}$$
(A.3)

Next, we consider the asymptotic variance matrix of $\hat{\pi}$ which we denote by $V(\hat{\pi})$. Note that, asymptotically, the set of the $\hat{\gamma}(\underline{i}')$ is independent of the set of $\hat{\gamma}(\underline{i}|\underline{i}')$. A similar statement applies to any distinct vectors $\chi(\underline{i}')$, $\chi(\underline{j}')$, $\underline{i}' \neq \underline{j}'$. This is clear from the block diagonal information matrix which is easily obtained from F. Linearizing the $\hat{\pi}(\underline{i})$ by a Taylor approximation around the $\gamma(\underline{i}')$ and $\gamma(\underline{i}|\underline{i}')$, we obtain (for large samples)

On letting

$$V(\hat{\chi}) = \frac{1}{n_1 + n_2} \left[D(\chi) - \chi \chi' \right] \equiv ((v_{m,n})) \quad m, n = 1, \dots, k'$$

$$V\left[\hat{\chi}(\underline{j}')\right] = \frac{1}{n_2 \gamma(\underline{j}')} \left[D(\chi(\underline{j}')) - \chi(\underline{j}')\chi'(\underline{j}') \right] \equiv V_{\underline{j}}, \quad (A.5)$$

where $D(\cdot)$ is a diagonal matrix with the vector (\cdot) on the main diagonal, we have asymptotically

$$V(\hat{\pi}) = \sum_{m=1}^{k'} \sum_{n=1}^{k'} v_{m,n^{\gamma}(\underline{j}_{m}^{\prime})} \chi(\underline{j}_{n}^{\prime}) + \sum_{\underline{j}^{\prime}} \gamma^{2}(\underline{j}^{\prime}) V_{\underline{j}^{\prime}} . \qquad (A.6)$$

When consistent estimators from (A.2) are substituted for the $\gamma(\underline{i}')$ and the $\gamma(\underline{i}')$ in (A.6), one obtains a consistent estimator $\hat{V}(\hat{\pi})$ for the dispersion matrix of the vector $\hat{\pi}$.

A Maximum Likelihood test of fit (i.e., $\gamma(\cdot, \underline{i}') = \gamma(\underline{i}')$ for all \underline{i}') is rather straightforward. The unrestricted MLE's are given by

$$\hat{\tilde{\gamma}}(\underline{i}') = \frac{n(\underline{i}')}{n_{1}}$$

$$\hat{\tilde{\gamma}}(\underline{i},\underline{i}') = \frac{n(\underline{i},\underline{i}')}{n_{2}}$$
(A.7)

Under the null hypothesis, (i.e., $\gamma(\cdot, \underline{i}') = \gamma(\underline{i}')$ for all \underline{i}') the MLE's of the $\gamma(\underline{i}, \underline{i}')$ are $\hat{\gamma}(\underline{i}, \underline{i}') = \hat{\gamma}(\underline{i} | \underline{i}') \hat{\gamma}(\underline{i}')$. On denoting the Maximum Likelihood Ratio (MLR) statistic by L, we have

$$-2 \log L = -2 \left\{ \sum_{\underline{j}'} n(\underline{j}') \log \left[\frac{\widehat{\gamma}(\underline{j}')}{\widehat{\gamma}(\underline{j}')} \right] + \sum_{\underline{j}} \sum_{\underline{j}'} n(\underline{j},\underline{j}') \log \left[\frac{\widehat{\widehat{\gamma}}(\underline{j},\underline{j}')}{\widehat{\gamma}(\underline{j},\underline{j}')} \right] \right\}. \quad (A.8)$$

Under the null hypothesis, this is asymptotically distributed as a central Chi-square variate with (k'-1) d.f.

Often, having established the fit, the experimenter will be interested in further inference on π based on the efficient estimator $\hat{\pi}$. In most practical problems, it is not feasible to obtain simple MLE's

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of π under further functional restrictions on the $\pi(\underline{i})$ (given that the model fits). One can verify this by trying to obtain the MLR test for independence in a 2 × 2 table. Even for this simple problem, the MLE cannot be obtained explicitly and one must call upon numerical techniques. In general, the usual log-linear hypotheses on π (hypotheses such as $C\pi = Q$ or $C[\log(\pi)] = Q$, where C is a contrast matrix, i.e., $C'\underline{1} = \underline{0}$) will impose complicated functional relationships among the $\gamma(\underline{j}|\underline{j}')$ and the $\gamma(\underline{j}')$. The MLE's will need to be obtained by some numerical computer subroutines.

The practical approach is to utilize the estimator $\hat{\pi}$ and the consistent estimator of its variance matrix, $\hat{V}(\hat{\pi})$, as initial input to the asymptotically equivalent least squares procedures presented in Grizzle <u>et al.</u> (1969) and Forthofer and Koch (1973). This is discussed in greater detail in the final section of this appendix where a convenient technique is given for implementing the Maximum Likelihood approach at the first stage and then proceeding with the Weighted Least Squares approach in the final stage using a single computer program.

Inference Based on Least Squares Estimators (LSE) of the $\pi(i)$

Before discussing the Least Squares approach (which will resemble to some extent that in Koch <u>et al.</u>, 1972), additional notation is required. Let $p(\underline{i}') = n(\underline{i}')/n_1$, $p(\underline{j},\underline{i}') = n(\underline{j},\underline{j}')/n_2$, and \underline{p}_1 be the vector whose elements are all $p(\underline{j}')$. Similarly, let \underline{p}_2 be the vector of length k'k obtained by stretching out all the $p(\underline{j},\underline{j}')$ in order. Finally, let $\underline{\gamma}_1 = E(\underline{p}_1)$, i = 1,2, and denote $\underline{p} = (\underline{p}_1',\underline{p}_2')'$.

The dispersion matrix of p is a block diagonal matrix V(p) with $V(p_i)$ on the diagonal, where

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$$V(\underline{p}_i) = \frac{1}{n_i} \left[D(\underline{\gamma}_i) - \underline{\gamma}_i \underline{\gamma}_i' \right] \qquad i = 1,2 .$$
 (A.9)

Next let F = Ap, where

$$A = \begin{bmatrix} A_{1}:(k'-1) \times k' & O \\ O & A_{2}:(k'k-1) \times k'k \end{bmatrix}$$
(A.10)

where A_1 is obtained from an identity matrix of dimension k' by deleting the last row and A_2 is similarly obtained from an identity matrix of dimension k'k. We can now write a model

$$E(\underline{F}) = X\underline{\theta} \qquad \underline{\theta}: (k'k-1) \times 1 . \qquad (A.11)$$

where $X = (X_1', X_2')'$ with X_2 being an identity matrix of order k'k-1 and X_1 of dimension $(k'-1) \times (k'k-1)$ has the form $X_1 = [I \otimes 1'|0]$ where I is the identity matrix of order k'-1, 1' = (1, ..., 1) and \otimes denotes Kronecker's product.

The variance matrix of \underline{F} is consistently estimated by $\hat{V}(\underline{F}) = A\hat{V}(\underline{p})A'$ where $\hat{V}(\underline{p})$ is obtained by substituting the unrestricted MLE's of the \underline{V}_i in the expression for $V(\underline{p})$. Thus, in large samples, one may use Weighted Least Squares to estimate the vector $\underline{\theta}$. The asymptotic LSE of $\underline{\theta}$ (which is BAN if (A.11) holds) is given by

$$\hat{\theta} = (X'\hat{V}^{-1}(\underline{F})X)^{-1}X'\hat{V}^{-1}(\underline{F})\underline{F}$$
(A.12)

and the consistent estimator of its dispersion matrix $V(\hat{\theta})$ is given by

$$\hat{\mathbf{V}}(\hat{\boldsymbol{\theta}}) = (\mathbf{X}^{\dagger} \hat{\mathbf{V}}^{-1}(\boldsymbol{F}) \mathbf{X})^{-1} \quad . \tag{A.13}$$

-43-

A test for goodness of fit is based on

$$X^{2} = \underbrace{F}^{\prime} \widehat{V}^{-1}(\underbrace{F}) \underbrace{F}_{} - \widehat{\theta}^{\prime} X^{\prime} \widehat{V}^{-1}(\underbrace{F}) X \widehat{\theta}$$
(A.14)

which, under the hypothesis that the model fits, follows an asymptotic Chi-square distribution with (k'-1) d.f.

If the model adequately describes the data, tests of hypotheses with respect to the parameters comprising $\hat{\theta}$ can be undertaken. Note that the elements of $\hat{\theta}$ are the k'k-l upper-left elements among the k'k parameters $\gamma(\underline{i},\underline{i'})$. The last element is obtained from the relation $\Sigma_{\underline{i}}\Sigma_{\underline{i'}}\gamma(\underline{j},\underline{j'}) = 1$. From $\hat{\hat{\theta}}$ and its estimated variance matrix, one can easily obtain the LSE $(\hat{\pi})$ of π and its estimated variance matrix $\hat{V}(\hat{\pi})$.

Employing the Maximum Likelihood Approach

Here we use notations from both of the two previous sections. The MLE's of the $\pi(\underline{i})$ and their asymptotic variance matrix have already been given. The overall procedure of first obtaining MLE's and then using asymptotic Least Squares theory appears somewhat inconvenient, especially when considering the available computer programs. Here, we discuss a simple technique to implement the MLE methodology, which can be employed using a single computer program. This approach is based on the fact that the MLE's of the $\pi(\underline{i})$ can be expressed as compound exponential-logarithmic-linear functions (see Forthofer and Koch, 1973) of the elements of p.

Specifically, we can write $\hat{\pi}$ (the MLE of π) as

$$\hat{\pi} = Q\left\{ \exp[K\log(A_p)] \right\}$$
(A.15)

$$(2+k)k' \stackrel{A}{\times} (k+1)k' = \begin{bmatrix} a_1 I:k' \times k' & a_2 I \otimes j':k' \times k'k \\ 0:k' \times k' & I \otimes j:k' \times k'k \\ 0:k'k \times k' & I:k'k \times k'k \end{bmatrix}$$

$$a_i = n_i/(n_1 + n_2), i = 1,2$$

where

where the unspecified identity matrix I has dimension k' and <u>l</u> is of length k.

$$Q = I \otimes 1'$$
.
 $k \times k'k$

Thus, on letting $\underline{y} = A\underline{p}$ and $\underline{z} = \underline{exp}[Klog(\underline{y})]$, we can conveniently write the asymptotic variance matrix of $\hat{\pi}$ as

$$V(\hat{\pi}) = QD(z)KD^{-1}(y)AV(p)A'D^{-1}(y)K'D(z)Q', \qquad (A.16)$$

where D(y) is a diagonal matrix with the vector y on the main diagonal.

As noted earlier, the vector $\hat{\pi}$ and the estimated variance matrix $\hat{V}(\hat{\pi})$ (which is obtained by substituting $\hat{V}(\underline{p})$ for $V(\underline{p})$) are subsequently used as initial inputs for further modeling based on Weighted (asymptotic) Least Squares procedures as in Grizzle <u>et al</u>. (1969). Thus, one may obtain functions of the $\hat{\pi}$ which are of interest for further modeling via a repeated chain of linear, log or exponential transformations, and then express a linear model for the resulting functions. The model can

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be tested for fit; given an adequate fit, linear hypotheses regarding its estimable parameters can be tested. And this entire procedure can be carried out in a single computer run using the new program GENCAT given in Landis and Stanish (1975).

APPENDIX B

The Supplementary (Telephone and Hospital) Data

| Po Po | lice | | | 16 | 5-55 | | | | | | 56 | 5+ | | | |
|---------------|---------|------|---------|----|------|--------|----|-------|-----|--------|-----|----|--------|----|-------|
| Non- | 1.00 | No | o Injui | ry | | Injury | , | | No | o Inju | iry | | Injury | | |
| Police | | U | L | LS | U | L | LS | lotal | U | L | LS | U | L | LS | lotal |
| iry ne) | n | 1035 | 12 | 1 | 24 | 1 | 1 | 1074 | 159 | 2 | 0 | 4 | 1 | 0 | 166 |
| Inju epho | | 122 | 122 | 14 | 5 | 3 | 0 | 266 | 21 | 11 | .1 | 0 | 0 | 0 | 33 |
| No (Tel | ΓS | 63 | 24 | 69 | 0 | 1 | 1 | ·158 | 10 | 6 | 8 | 1 | 0 | 0 | 25 |
| y (1) | ∍ | 210 | 4 | 0 | 279 | 3 | 2 | 498 | 17 | 1 | 0 | 27 | 0 | 0 | 45 |
| njur. spit | | 26 | 19 | 3 | 8 | 18 | 4 | 78 | 2 | 0 | 0 | 3 | 3 | 0 | 8 |
| II Hos | ΓS | 6 | 1 | 9 | 5 | 3 | 9 | 33 | 3 | 0 | ١ | 0 | 0 | 0 | 4 |
| Tota | 1 | 1462 | 182 | 96 | 321 | 29 | 17 | 2107 | 212 | 20 | 10 | 35 | 4 | 0 | 281 |

Table B.1 Belt usage by injury level for supplementary sample, controlling for age.

No. missing observations = 13

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| Pc | olice | | | Ma | le | | | | | | Fem | ale | | | |
|---------------|--------------|------|--------|----|-----|--------|----|-------|-----|--------|-----|-----|--------|----|-------|
| Non- | | No |) Inju | ∽у |] | Injury | | | No | o Inju | ry | | Injury | | |
| Police | \mathbf{n} | U | Ļ | LS | U | L | LS | Total | U | L | LS | U | .L | LS | Total |
| iry ne) | D | 813 | 10 | 1 | 18 | 1 | 0 | 843 | 381 | 4 | 0 | 9 | 0 | 1 | 395 |
| Inju lepho | | 106 | 102 | 12 | 3 | 3 | 0 | 226 | 36 | 32 | 3 | 2 | 0 | 0 | 73 |
| No (Tel | ΓS | 57 | 24 | 59 | ۱ | 1 | 0 | 142 | 16 | 6 | 18 | 0 | 1 | ١ | 42 |
| y al) | Ð | 121 | 3 | 0 | 173 | 2 | 2 | 301 | 106 | 2 | 0 | 133 | 2 | 0 | 243 |
| njur spit | | 19 | 12 | 2 | 8 | 14 | 3 | 58 | 10 | 7 | 1 | 2 | 7 | 1 | 28 |
| I H) | LS | 6 | 0 | 6 | 3 | 3 | 6 | 24 | 3 | 1 | 4 , | 2 | 0 | 3 | 13 |
| Total | | 1122 | 151 | 80 | 206 | 24 | 11 | 1594 | 552 | 52 | 26 | 148 | 10 | 6 | 794 |

Table B.2 Belt usage by injury level for supplementary sample, controlling for sex.

No. missing observations = 13

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| Po | lice | | | U. | .s. | | | | | | For | eign | | | |
|--------------|-----------------|------|--------|-----|-----|--------|----|-------|-----|--------|-----|------|--------|---------------------------|-------|
| Non- | | No | o Inju | ry | | Injury | | | No | o Inju | ry | | Injury | · · · · · · · · · · · · · | |
| Police | $\overline{\ }$ | U | L | LS | U | L | LS | Total | U | L | LS | U | L . | LS | Total |
| ry ne) | ח | 1086 | 13 | 0 | 23 | 2 | 1 | 1125 | 101 | 1 | 1 | 5 | 0 | 0 | 108 |
| Inju epho | | 130 | 120 | 14 | 5 | 3 | 0 | 272 | 11 | 13 | 1 | l 1 | 0 | 0 | 26 |
| No (Tel | LS | 52 | 24 | 55 | 1 | 0 | 1 | 133 | 21 | 6 | 21 | 0 | 1 | 0 | 49 |
| y (la | Ð | 197 | 4 | 0 - | 273 | 3 | 1 | 478 | 29 | 1 | 0 | 32 | 0 | 1 | 63 |
| njur spit | 1 | 26 | 18 | 1 | 7 | 20 | 3 | 75 | 2 | 1 | 2 | 3 | 1 | 1 - | 10 |
| I I Ho | ΓS | 7 | 1 | 6 | 3 | 3 | 7 | 27 | 2 | 0 | 4 | 2 | 0 | 2 | 10 |
| Total | | 1498 | 180 | 76 | 312 | 31 | 13 | 2110 | 166 | 22 | 29 | 43 | 2 | 4 | 266 |

Table B.3 Belt usage by injury level for supplementary sample, controlling for "make" of car.

No. missing observations = 26

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| Po | lice | | | 1960 | -1968 | | | | | <u></u> | 1969- | -1971 | · | | |
|--------------|-------------|-----|------|------|-------|-------|----|-------|-----|---------|-------|-------|-------|----|-------|
| Non- | 1100 | Nc | Inju | ry | | Injur | у | | No | o Inju | ry | | Injur | у | |
| Police | \setminus | U | L | LS | U | L | LS | Total | U | L | LS | U | L | LS | Total |
| ury one) | U | 399 | 1 | 0 | 10 | 0 | 0 | 410 | 368 | 5 | 1 | 6 | 1 | 1 | 382 |
| Inj eph | | 46 | 21 | 0 | 3 | 1 | 0 | 71 | 36 | 40 | 3 | 2 | 1 | 0 | 82 |
| No Tel | ΓS | 5 | 1 | 0 | 0 | 0 | 0 | 6 | 14 | 9 | 9 | 1 | 0 | 0 | 33 |
| y (1) | n | 84 | 2 | 0 | 117 | 1 | 0 | 204 | 64 | 2 | 0 | 87 | 1 | 0 | 154 |
| njur spit | | 8 | 5 | 1 | 4 | 5 | 0 | 23 | 7 | 7 | 1 | 2 | 8 | 0 | 25 |
| I (Но: | ΓS | 0 | 0 | 0 | 1 | 0 | 0 | 1 | Ō | 0 | 0 | 2 | 1 | 1 | 4 |
| Tota | 1 | 542 | 30 | 1 | 135 | 7 | 0 | 715 | 489 | 63 | 14 | 100 | 12 | 2 | 680 |

Table B.4 Belt usage by injury level for supplementary sample, controlling for model year.

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| Police | raan in tee | u na pornegne i sun | 1972 | -1973 | ~ % | | | | | 1974 | -1,975 | | - 1 . | |
|--------------------|----------------|---------------------|--------|-----------|-------------|-----|-------|------|-------|----------|-----------|------------|------------------|------------------|
| Non- | ldiviti N | o Inju | су | 2 . 2 | Inju | ry | т | N | o Inj | ury | х т ум | Injur | y | T 1 1 |
| Police | U tr | L | LS | U | L | LS | IOTAI | U, | Ļ | LS | IJ | ĻL | LS | |
| ury ur U | 296 | .4 | .0 | 8 | .: 0 | . 0 | 308 | 129 | 5 | .0 | .3 | -1 | 0 | 138 |
| Inji léphi L | 51 | 58 | 8 | | ٦ | 0 | 118 | × .9 | 13 | 2 | ы | 0 | 0 | 25 |
| LS LS | · 10 | ••• •6 | 11 * 1 | · · 0 · | - 0 | - 0 | 27 | 43 · | 14 | 57 | 0 | 1 | Ĩ | 116 |
| ý al) | 54 | 0 | 0 | 60 | [] | 1 | 116 | -25 | 1 | 0 | 35 | .0 | ·] | 62 |
| njúr spít | 11 | 6 | .] | | .7 | 2 | 30 | 2 | 1 | <u>0</u> | 1 | 2 | .2 | 8 |
| HOH) ST | [~] 2 | 1 | 1 | · 1 | 1 | 0 | 6 | 7 | 0 | 9 | 1 | 1 | 8 | 26 |
| Total | 424 | 75 | 21 | 72 - | 10 | - 3 | 605 | 215 | 34 | 68 | 41 | 5 | 12 | 375 ⁻ |
| | | | · | ÷ · · | | | | - | | | - | <u>.</u> . | ···· ·· · | |

Table B.4 Continued.

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No. missing observations = 24

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| Po | lice | | | Mir | ior | | | | | | Sev | ere | | | |
|---------------|--------------|-----|-------------|-----|-----|-------|----|-------|-----|--------|-----|-----|--------|----|-------|
| Non- | | No | Inju | ŗy | I | njury | | | N | o Inju | iry | | Injury | , | |
| Police | \mathbf{n} | U | L | LS | U | L | LS | Total | U | L | LS | U | L | LS | Total |
| ry ne) | U | 613 | 6 | 0 | 9 | 0 | 0 | 628 | 401 | 6 | 1 | 17 | 1 | 0 | 426 |
| Inju epho | | 79 | 64 , | 6 | 1 | 1 | 0 | 151 | 27 | 40 | 6 | -2 | 2 | 0 | 77 |
| No (Tel | LS | 35 | 14 | 33 | 1 | 0 | 1 | 84 | 24 | 11 | 33 | 0 | 1 | 0 | 69 |
| بر الق | n | 82 |] | 0 | 61 | 2 | 0 | 146 | 112 | 2 | 0 | 197 | 1 | 1 | 313 |
| njur spita | _ | 12 | 7 | 0 | 4 | 2 | 2 | . 27. | 12 | 8 | 2 | 6 | 15 | 1 | 44 |
| I SoH) | ΓS | - 3 | 0 | 4 | 1 | 1 | 3 | 12 | 5 | Ŋ | 6 | 3 | 2 | 6 | 23 |
| Total | | 824 | 92 | 43 | 77 | 6 | 6 | 1048 | 581 | 68 | 48 | 225 | 22 | 8 | 952 |

Table B.5 Belt usage by injury level for supplementary sample, controlling for vehicle damage severity (TAD).*

* Minor = FR-LF,1; OTHER, 1-3

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Severe = FR-LF, 2-7; OTHER, 4-7

No.missing observations = 399

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Table B.6 Belt usage by injury level for supple-mentary sample, controlling for accident type.

| | lica | | | _Co1 | lision | | | | • • • • • • • | | Non-Co | llisio | n . | سیم میں . د | |
|-------------------|--------------|------|-------------|------|--------|-------|-----|-------|---------------|----------|--------|------------|--------|----------------|-------|
| Non- | | No | <u>Inju</u> | ry |] | njury | | | Ņc |) Inju | ry j | | Injury | | |
| Police | | U | L, | ĻŞ | U, | Ļ | LS | Total | U | L | LS | U | μ, | LS | Total |
| iry ne) | Ð | 988 | 11 | 1 | 1,8 | 1 | 1 | 1020 | 197 | 3 | 0 | 1:0 | . 1 | 0 | 211 |
| Inju | - - | 120 | 113 | 10 | 1 | 1 | 0 | 245 | <i>:</i> 20 | 20 | 4 | . ; 4 | 1 | 0 | 49 |
| No (Te | ۲S | 65 | 23 | 56 | ٦ ا | 1 | 0 | 146 | 6 | 7 | 20 | 0 | 0 | 1 | 34 |
| (1) | Ū. | 185 | · 4 | 0 | 19.8 | 3 | 1 | 391 | 41 | -1 | 0 | 108 | 0 | ·J | 151 |
| nju ry sp'i ta | : - است. | 25 | 16 | 1 | 9 | 14 | 3 | - 68 | : 3 | .3 | 2 |] | 7 | 1 | 17 |
| (Ho: | ۲S | 5 | 1 | .6 | 4 | 1 | . 7 | 24 | 3 | 0 | 3 | _ 1 | 1 | 2 | 10 |
| Total | с - 1 | 1388 | 168 | 74 | 231 | 21 | 12 | 1894 | 270 | 34 | 29 | 124 | 10 | 5 | 472 |

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No.missing observations = 30

APPENDIX C

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Accident Report Information Standard North Carolina Accident Report Form The following variables from the police accident report form (shown in this Appendix) were utilized in the analysis:

- 1. Vehicle # (as assigned by police agency)
- 2. Month of Accident (January June)
- 3. Day of Week
- 4. Hour of Day (e.g., 8:00 8:59 a.m.)
- 5. <u>Accident Type</u> (e.g., collision of motor vehicle in road with pedestrian)
- 6. Driver's (or Injured Passenger's) Year of Birth (e.g., 1952)
- 7. Driver's (or Injured Passenger's) Sex
- 8. Driver's (or Injured Passenger's) Race (white, non-white)
- 9. Vehicle Year (e.g., 1971)
- 10. Vehicle Make (e.g., Plymouth)
- 11. <u>Vehicle Type</u> (e.g., two or four door sedan (passenger vehicle), stations wagon (passenger))
- 12. V.I.N. (Vehicle Identification Number)
- 13. <u>First TAD</u> (location and severity; e.g., FD3 = front distributed of relative severity 3)
- 14. <u>Police Reported Injury</u> (Injuries for driver or injured passenger)
 - K Killed
 - A Serious injury
 - B Moderate injury
 - C Minor injury
 - 0 No injury
- 15. <u>Restraint Used</u> (the individual being coded) as Reported by Police

None Lap Belt Shoulder Belt Only Lap and Shoulder Belt Child Restraint System Not Recorded

| 1. LOCALITY 1. Business 2. Residential 3. School & pierground 4. Open country 3. BFED LUNT 3. RFGD LUNT 3. RFGD LUNT 3. RIAD FEATURE 3. Alley intersection 4. Intersection of two Monitories 5. Monitories 1. Concerner 1. Concerner 3. Concern | ROAD DEFECTS Lease meterial on surface Malas, deep vuts Law shaulders Satishaulders Other defects Road under construction No defacts ROAD CONDITION Dorder Snawy Licki T CONDITION Derhouss (street lighted) Derhouss (street instruct | WEATHER Clower Clower Clower Clower Clower Clower Staning Snawing Snawing Snawing Stast at hell TRAFFIC CONTROL Step and ge signel Step and ge signel Step and ge signel Step and ge signel R. R. soft signel fisher R. R. soft signel fisher R. R. Soft signel fisher Notice device Other Other device Na control greacent De DECT STRUCK (liter Tree only) | Guerdrail ar guerdpart Is medica Guerdrail ar guerdpast en shoulder Bridge Underpass Troffic island, surb, ar medica Sign ar sign post Article obscience Anter object Pedestrian Pedestrian Bridge object Orinking-shilling inparted | 3. A sloop 4. Other physical logatmast 5. Restriction net complied with 6. Narmal 7. Condition net hours 13. CheditCAL TEST 14. PEDESTAINA ACTION 1. Crossing at intersection 3. Consing from behind particle solution 1. Consing from behind particle solution 5. Walking with restle 5. Walking earlier traffic 6. Getting on at other 1. Stoplay in read 8. Philing in read 8. Philing in read 10. Units 11. Other 12. Not in read 15. VENICLE MANEUVER 1. Stoplay in Particle Inc. | 3. Parked in travel lance 4. Going wright shead 5. Chenging into ar marging 4. Passing 7. Making right turn 8. Making bif turn 9. Making U turn 10. Beeking 11. Storing ar stopping 12. Storing in readway 13. Parking 14. Leaving parked pasition 15. Other 16. VEHICLE DEPECTS (List and a marg) 1. Defaction backlights 3. Defactive strange 5. Defactive strange |
|--|---|--|---|---|---|
| Date of Accident | Day 19 Week | of | A.M. P.M. Hour | | Do not write in this space |

| | NOL | Occur | red | | | County | | City or Town of _ | | | | | _ | | | | |
|----------|--|--|-------------------------------|-------------|-----------------|--|----------------------|---------------------------------------|--|--|------------------|-----------------|--------------|------------|--------------------------------|--|-----------|
| | Š. | Outsi | de C | ty or | Town . | | ᇢ뎥 | çç•' | Lim Lim | • | Cente | r | L | | | 1 | |
| | 121- | 0n | | | | | | | | | | | | | | Patrol Area | |
| | | | C |) Mile | , ri wj | · No. (I., U.S., | At or | ·., R.U.) | to., or with | in corpora | e lim | its, i | dentity | by | name | L | |
| | - | (0 F. | - 5 |] Fee | . 5 | | From | Hwy. No., o | Adiacent | County Li | | <u> </u> | Towar | ۰ <u> </u> | Hwy. No., City, pr. | Adiacent County L | ine |
| | | T | | Ra | n off f | Rood | Non-C | ollision in R | 0 ad | | | - 6 | allisia | n of | Motor Vehicle in Road | With: | |
| | Ξ | 1.1 | Right | 2. L | 1 3. | Straight Ahead | 4 Overt | um 5. Other | in Road 6 | Pedestria | 7. F | arks | d Vehi | | - Train 9. Bicycle 10 | Animal 11. Fixed | 12.0 the |
| | <u> </u> | : | | 1 | | | | | | | 1. | | | | | Оы. | |
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| | | 13 | . Rea | r End | 14, R | ear End 15.L | eft Turn Rondwovi | 16. Left Turr Cross Troffi | 17. Right | Turn 18 ndwny C | Righ | tt Tu Troffi | rn 19. | Hea | d On 20. Sideswipe 2 | 1. Angle 22. Boo | :king |
| | | | VE | ICL F | NO. 1 | | Nous non | | | | | VE | ICLE | NO. | 2 or PEDESTRIAN | | |
| | No. | of | | | | | | | | | | | | | | | |
| | Invo | lved | Dri | ver: | First | M | i ddl e | | Last Nan | Drive: | " | irst | | | Middle | La | at Name |
| | | | | | | | | | | | | | | | | | |
| | <u> </u> | | ~ 0 | ress: | | | | States | | | | | | | | . Santas | |
| | C.1.9 | | | | | | Y | as No | | | | | | | D : | Y++ No | |
| | 18 0 Ros | 100 ve | 86876 ., | 155 SCI | ne as Núver | on Driver's Lic | ense? | J L. | - | Baca | sve d | ddres | s som 0 | | on Univer's Licenser In Lic | LJ LJ ••••• | |
| F I | Dat | a of P | irth: | ' |)ri ve r | s cic:S | acify Res | triction: | | Date | v Jen: Sf Rij | | | | Specify R | atriction: | |
| | | | | Mont | D | ay Year | | | | | | M | on th | De | y Year | | |
| Ĕ | Mem | ber of | 1 | Yes | No. | Veh. V | eh. | Ý | nh. | Membe | r of | .) | <u>(es N</u> | • | Veh. Veh. | Veh. | |
| Ē, | Lie | ea ra Píate | No | <u> </u> | ب | 1 ear: N | State: | T | | | core ate l | ₩∓ (Νο | | ر | State | Ype: | |
| 8 | VIN | | | | | | ODOM | | | VIN | | | | | OD | OM, | |
| ¥ | 0- | •r: | | | | | | | | Owner | _ | | | | | | |
| P F | Add | ress: . | | | | | | | | Addres | . s: | | | | | | |
| | City | • | _ | | | | | | | _ City:_ | | | | | | State: | · |
| | Part | hs. | | | | | | Amount | | Parts | | | | | | Amount | |
| | Dam | aged (| TAD |) | | | | f Domage \$ | | |) bei | TAD) | | | | of Damage \$ | |
| 1 | DAV | aple: | Va | hicla | | | | | | Dri vol | de: | ـ | -1.0 | | | | |
| | T • • | Nº. | Ren | no ved | to: | | | | | _ "" | Nº. | Remo | ved to: | | | | |
| | | | | | | | | | | | | | | | | | |
| | By: | | _ | | | [/] | luthority: . | | | - By: - | | | | | Authori | ty : | |
| | Othe | er . | | | | | | Am | t. of Dam. | Owner | an d | | | | | | |
| | P rop | erty I | Jamo | ged | _ | | | (> | DV CECT | | UCTI | ONS | | | | | |
| a | Give | iniury | clas | . restr | aint us | ed, rare sex an | d ana of a | il occupants i | n the snace | CORRESON | | | | | a Mamaa and addaaaaa | | |
| | were | injuri | id. Fo | r type | of Re | traint (Res.) use | d: N - Non | e, L — Lap Be | It, LS Lap | and Should | ser, S | Sh | oulder | Belt c | only, YR - Child Restrain | are necessary for pe it System. | rions who |
| Ż | K = K | illed | A= | Incop | aci toti | ng B=Noninc | ap aci tatin | g - Injury oth | er than K a | r A eviden | 1 01 1 | he so | | =Ng | visible sign of injury | but complaint 0 | =No inju |
| <u>é</u> | | | <u> </u> | | - | | | | | | | 1 | | 10 | pain, momentary unco | nsciousness | |
| ž | SEAT | lei | Res | Race | ^9 • | INJURE | DNAMES. | AND ADDRE | SSES | SEAT | in) cl | Res | Race | Age | INJURED NAME | S AND ADDRESS | ES |
| Ż | | | <u> </u> | | | First Nome | | L. | 247 | | <u> </u> | - | - | _ | FirstName | Last | |
| ā | Left | | | | | | DRIVE | R I | | Left | | | 1 | | DRIVER 2 | OR PEDESTRIA | N |
| | Front | ╋ | - | | ┝╧┥ | | | | | - Pront | | {— | | <u> </u> | | | |
| | Cente | | | | † | | | | | | | 1 | | | | | |
| | Front | | [] | | | | | | | Front | | | | | | | |
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| 71 | Right | | | | l r | | | | | I FINAT | | + | ł | | | | |
| Ē | Right Front | \downarrow | - | | ┝╌┦ | | | | | | | 1 | | | | | |
| i i i | Right Front | - | | | ┝╌┨ | | | · · · · · · · · · · · · · · · · · · · | | | | 1 | | | | | |
| (n i i | Right Front Left Rear | | | | | ······································ | | · · · · · · · · · · · · · · · · · · · | | Left Rear | | | | | | | |
| (init) | Right Front Left Rear | | | | | | | | | Left Regr | | - | | | | | |
| (init) | Right Front Left Rear Cente | - | | | | | | | | | | | | | | | |
| (Init) | Right Front Left Rear Cente Rear | | | | | | | | · · · · · · · · · · · · · · · · · · · | Left Rear Center Rear | | | | | | | |
| | Right Front Left Rear Cente Rear | - | | - | | | | | | Left Rear Center Rear Right | | | | | | | |
| [1]4]) | Right Front Left Rear Cente Rear Right Rear | | | | | | | | | Center Rear Center Rear Right Rear | | | | | | | |
| (late) | Right Front Rear Cente Rear Right Rear Total | r No. 0 |) cour | en 13 | | | Toral No. | lni. | | Center Rear Center Rear Right Reor Total | No. 0 | Coup | ants | | Totel A | 10. lnį. | |
| (Init) | Right Front Rear Cente Rear Right Rear Total Injura | r No.C | en to | onts: | | | Toral No. | lni. | | Center Rear Center Rear Right Reor Total | No. 0 | Ccup | ants | | Totel N | lo. Inj. | |
| | Right Front Rear Cente Rear Right Rear Tatal Injuri WiT- | r No.C | | ion ts : | | | Torel No. | lni. | Address | Center Rear Center Rear Right Reor Total | No. 0 |)c cup | onts | | Totel A | Io. Inj. | |
| | Right Front Rear Cente Rear Right Rear Total fnjurn WIT- NESS Arres | r no. C od tak Na iES Ni | | ion ta | | | Total No. | lnį | Address . Address . Charce(s) | Center Rear Right Reor Total | No. 0 |)ccup | onts | | Total M | lo. Inj. | |
| | Right Front Rear Cente Rear Right Rear Tatal Injuri WIT- NESS Arres | r nd tak Na iES Na is: Na |) coup en to ame ame | on ts | | | Total No. | .inj. | Address . Address . Charge(s) | Center Rear Center Rear Right Reor Total | No. 0 | 2ccup | gn ta | | To tol A Ph Ph (Ci | Io. Inj. one No one No one No | |
| | Right Front Rear Cente Rear Right Rear Tatal fnjurn WIT- NESS Arres Sign | r ed tak Na ES Ni Na Na Na Na |) c cup en fo me me | on ta | | | Total No. | ini. | Address - Address - Cherge(s) Cherge(s) | Left Rear Center Rear Right Reor Total | No. 0 | Coup | gnts | | To tol 4 | Io. Inj. one No one No 1. No.) i. No.) | |









| Vehicle VIOLATION INDICATED | EMERGENCY ASSIS | TANCE | RESERVE | D FOR STATE U | SE: | • | |
|---------------------------------|-----------------|---------|---------|---------------|------------|---------------------------------------|--------|
| 1 2 | INFORMATION | 4 | 20. | 21. | 22 | 23. | 24. |
| 2. Excessive Speed | INVESTIGATOR | [] am | 25 | 26. | 27. | 28. | 29. |
| [] [] 3. Yield Violation | NOTIFIED | p.m. | RESERVE | D FOR CITY OR | OTHER USE: | | |
| C 4 Left of Center | 8Y | | | | | | |
| S. Passing Violation | | | | | | | |
| 🛄 🛄 6. Stop 5. or Yield S. Vio. | INVESTIGATOR | 🗋 a.m. | | | | | ····· |
| 7. Traffic Signal Via. | ARRIVED | 🗔 p.m. | | | | | ······ |
| 📑 🚍 8. Safe Movement Vio. | | | | . <u>.</u> | | | |
| - 9. Too Close | AMBULANCE | [] o.m. | | | | | |
| - 10. Improper Turn | ARRIVED | [] p.m. | | | | · · · · · · · · · · · · · · · · · · · | |
| [] [] 11. Improper or No Signal | OTHER COMMENTS: | | | | | | |
| [] 12 Improper Parking Location | | | | | | | ····· |
| 13. Other Improper Driving | | | | | ···· | | |
| (describe) | ļ | | | | | | |

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APPENDIX D

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Telephone Interview: Introduction Format, Questionnaire

TELEPHONE INTRODUCTION FORMAT

Hello M. ______, my name is ______ and I am with the University of North Carolina Highway Safety Research Center. The Department of Transportation in Washington, D.C. is continually trying to learn more about seat belt usage and corresponding effectiveness in reducing deaths and injuries in highway crashes. In this connection, we are doing a survey of North Carolina drivers who were recently involved in a traffic accident, primarily to find out how they feel about seat belts in general and whether the seat belt might have helped (or hindered) in the accident in question. Would you mind answering a few brief questions? Thank you.

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Note:

(1) If the person we need to talk with is <u>not at home</u>, try to find out when a good time to call back and reach him would be. Very generally explain that your name again is _______ and that you work for the University of North Carolina. As part of a telephone survey, you are calling people to find out about automobile seat belts and their usage.

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(2) If the interviewee wants to know more about HSRC, the explanation can be derived from the following:

> The University of North Carolina Highway Safety Research Center (HSRC) was created by a statute of the 1965 North Carolina General Assembly, and was directed by the Governor to perform three functions:

- evaluate North Carolina's existing highway safety programs.
 - coordinate and participate in the professional training of persons involved in highway safety.
 - close the gap between knowledge created by highway safety research and its use in saving lives.
- (3) If the person needs to know how we know about his accident, explain that all accidents are public record at the Department of Motor Vehicles (DMV) in Raleigh and being a research organization engaged in highway safety research, we often need to have access to these records.
- (4) If the individual seems upset, suggest that he feel free to call HSRC <u>collect</u> at (919) 933-2202 and ask for Dr. Campbell or Dr. Reinfurt for further information about this survey.

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NAME OF DRIVER

1. Was the vehicle you were in during your accident a passenger car (); truck ()? Do you know the make, model and approximate year of the car? _____ If not, was the car a large car (Olds, Buick), intermediate (Chevelle), or small (Vega). (IF TRUCK, YOU ARE FINISHED) 2. Does your car (the one in the accident) have seat belts? Yes () No () (IF NO, GO TO 5) () Lap () Lap and shoulder 3. If so, what kind of belts? () Don't know () Not sure about shoulder part Nere you wearing your seat belts? 4. Lap only () () No belt
() Shoulder only
() Unknown or don't remember Lap and shoulder () For those who were wearing their shoulder belt: Since you were wearing your shoulder belt, can you tell me if yours is the kind that allows you some freedom of movement while you're belted in? (If they need an explanation use turning on the radio or opening the glove compartment as illustrations of freedom of movement). Yes () No () Can't say or don't remember () In your accident did the shoulder belt hold you in place? In other words, did it "lock up" like it was supposed to? No () Can't say or don't remember () Yes () If they don't remember ask, "Did you feel like any part of your

chest had been bruised or was sore after the accident from where the shoulder belt went across your chest?" Yes () No () "Did your waist feel especially sore from the lap belt?" Yes () No ()

Did you hit the steering wheel at that time?

Yes () No () Can't say or don't remember ()

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- 5. Did the officer ask you if you were wearing a seat belt?
 - Yes () No () Can't say or don't remember ()

6. Were you injured? Yes () No ()

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If YES, would you describe your injury as slight (), moderate (), severe ()?

Could you please describe where your injuries were and what types of injuries you had.

7. If you weren't really injured, can you recall if you had any aches or pains?

Yes () No ()

If yes, can you describe where?

APPENDIX E

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Hospital Survey: The Hospital Report Form

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Highway Safety Research Center University of North Carolina Chapel Hill

AUTOMOBILE INJURY AND SEAT BELT DATA HSR-006

Instructions: Please complete one form for each patient treated for injuries due to an automobile crash. Return the form to HSRC in the attached pre-addressed envelope. No stamp is necessary. If you have any questions please feel free to call collect Ms. Lucy Smith or Ms. Jane Youngblood at (919) 933-2202.

| 1. | Patient's Name | | Middle | | |
|----|----------------------|------------------------|--|--------|----------|
| | Firs | τ | midale | | Last |
| | Date of Birth | | Date Trea | ited | |
| 2. | Patient's AddressS | treet or | P.O. Box | | |
| - | -0 | ity | St | ate | Zip Code |
| 3. | Safety Restraint Use | : a. b. c. d. | No Belt Lap Belt Only Both Lap and Shoulder Unknown | • Belt | |
| | · . | | | | |
| | · · · · | | | · | |

Name and Title of Person Completing Form:

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-67-ABBREVIATED INJURY SCALE



| SEVERITY CODE | SEVEPITY CATEGORY/IN JURY DESCRIPTION | |
|------------------|---|--|
| • | SERIOUS (Life-Threatening, Survival Probable) | |
| | GENERAL Severe lacerations and or avulsions with dangerous hemorrhage. 30-50% surface 2° or 3° burns. | |
| | HEAD AND NECK Cerebral injury with or without skull fracture, with uncon- sciousness of more than 15 minutes, with definite apparmal neurological signs; post-trainatic amnesia 3-12 hours. Compound skull fracture. | |
| | CHEST Dpen chest wounds; flait chest; pneumomediastinum; myocatdial contusion without circulatory embarrassment; pericatdial injuries | |
| | ABDOMINAL Minor laceration of intra-apdominal contents (to incluse ruptures spleen, kianey, and infuries to tail of pancress). Introperitaneal bladder rupture Avulsion of the genitals Thoracic and for London spine fractures with paraplesia | |
| | EMITIES Multiple clased long-pone fractures. Amputation of Limos. | |
| 5 | CRITICAL (Survival Uncentain) | |
| | <u>OSNERA:</u> Over 50°- body surrace 1° or 3° auros. | |
| | HEAD AND NECK Cerebral injusy with an withour skull fracture with unconsciousness of mare than 24 hours, point number a america more then 12 hours intractanial theoremage, sign of increases intractanist pre-ure decreasing statulations of consciousness, brack-reactioned and e0, pro- g childenite in blood pre-ure on programice public regulations Cerebral statulations of a moustaine egic. Cerebral statulations | |
| | <u>CHEST</u> Chest injuries with major recollatory embarrossmont inscenation of traches, hemomediast num, etc. Vartic Iscenation. T4ysourdial ruck e an optic in with currupts y embarrossment. | |
| | ASDIMAL Subvice, publics or revice location of intra-Joophing veise to or algory, exclorigingly, power or uneter. Ex ISE MITIES | |
| L | | |
| | FAIAL | |
| F | SEVERITY UNKNOWN | |

- * Developed by the American Medical Association Committee on Medical Aspects of Automative Safety, in cooperation with physicians representing medical specialties most involved in the
 - diagnosis, care and treatment of crash injuries, and General Motors Corporation.

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4. Overall Severity of Injuries: (see p.2 for injury scale)
(1) Minor
(5) Critical
(2) Moderate
(6) Fatal
(3) Severe, Not Life Threatening
(4) Serious, Life Threatening

- 5. For Each Injury:
 - a. Indicate the location of the injury by marking on the drawing below.
 - b. Write the <u>degree</u> (e.g. major, slight, compound, l-in. etc.) and <u>nature</u> (e.g. burise, laceration, abrasion, fracture, burn, internal injury, etc.) of this injury.

EXAMPLE locaration of flesh _____ simple fracture of thumb



